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The effect of contrast randomisation on the discrimination of changes in the slopes of the amplitude spectra of natural scenes

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Abstract

It has been suggested (Tadmor & Tolhurst, 1994) that the psychophysical task of discriminating changes in the slope of the amplitude spectrum of a complex image may be similar to detecting differences in the degree of blur. It has also been suggested that human observers may perform this discrimination by detecting changes in the effective contrast within *single* narrow spatial-frequency bands, rather than by detecting changes in the slope *per se* which would involve the use of contrast information across many different frequency bands. To distinguish between these two possibilities, we have developed an experiment where observers were asked to discriminate changes in the spectral slope while different amounts of random contrast variation were introduced, with the purpose of disrupting their performance. This disruptive effect was designed to be particularly manifest if the observer really was performing a single frequency-band contrast discrimination but to be unnoticeable if the observer was discriminating the change of slope *per se*. Our results imply that the observers do not usually detect changes in contrast in just one narrow spatial-frequency band when they discriminate changes in the slope of the amplitude spectrum. Rather, they must compare contrast between bands or, at least, they use contrast information from more than one band. However, for edge-enhanced (whitened) pictures, there is some evidence to suggest that observers rely on contrast changes in only a limited, low spatial-frequency band.

1. Introduction

There has been a resurgence of interest in Kretzmer's (1952) discovery that that much of the variety of images which the visual system may naturally encounter are actually described by a simple statistical relationship (Carlson 1978; Burton and Moorhead 1987; Field 1987). The statistical regularity of natural images is shown by a simple relationship, describing the amplitude of the frequency coefficients in their Fourier spectra

$$amplitude(f) \propto f^{-\alpha} \quad \text{Eqn.1}$$

where amplitude is averaged across all orientations in the spectrum, f is spatial frequency and α (also called the *slope parameter* since the relationship is a straight line on log-log co-ordinates) is greater than zero. The value of α , although it varies from image to image, has been found to be within a fairly narrow range (0.7 - 1.5) for achromatic images (Burton and Moorhead 1987, Tolhurst *et al* 1992) and across the visible spectrum (Párraga *et al* 1988a). This means that the degree of correlation in luminance between nearby points as a function of distance between the points is remarkably similar from one natural scene to the next. It has been suggested that the overall organisation of the visual system including the response properties of individual neurons might exploit any redundancies or correlations in natural scenes (Barlow 1961; Srinivasan *et al* 1982; Laughlin 1981, Field 1987, 1989, Attick 1992, Attick and Redlich 1992, Brelstaff and Troscianko 1992; Olshausen and Field 1997, van Hateren and van der Schaaf 1998, Tadmor and Tolhurst 2000).

Knill *et al* (1990) hypothesised that the human visual system should be tuned to best discriminate those scenes with the same α -value as the real world and, therefore, that visual discriminations should be favoured when images have natural statistics (cf Parraga *et al* 2000). Knill *et al* and Tadmor and Tolhurst (1994) measured the threshold of human observers for discriminating changes in the slope (α) of stimuli derived either from patterns of dots of random luminance or from digitised photographs of natural scenes. Tadmor and Tolhurst argued that discriminating changes in the slopes of amplitude spectra is analogous to discriminating changes in the degree of blur (Hamerly and Dvorjak 1981, Watt and Morgan 1983, Walsh and Charman 1988, Hess *et al* 1989). A greater value of α is roughly comparable to a greater degree of blur (Goodman 1968). There had been speculation as to whether blur discrimination relied upon detecting changes in the contrast in some limited part of the frequency spectrum of the image (e.g. Hess *et al* 1989) and

Tadmor and Tolhurst (1994) tried to model the α discrimination task in terms of contrast discrimination too.

The contrast of a complex image such as a natural scene is not simple to evaluate, but must be calculated in several band-limited spatial scales (Peli 1990). Contrast is evaluated in several spatial-frequency bands of about 1 octave bandwidth, and so local *band-limited* contrast, evaluated at a particular spatial scale, may be considered as representative of the activity of one spatial-frequency channel (Campbell and Robson 1968).

Tadmor and Tolhurst proposed that the discrimination of changes in α could be modelled as a discrimination of changes in local contrast *within one specific frequency band*. They argued that changing the slope α of the amplitude spectrum also causes this local contrast of the image to change. The magnitude of this change (at one or more spatial scales) agrees reasonably well with that required to discriminate changes in Michelson contrast of a simple sinusoidal grating of the appropriate frequency. The apparent success of this local band-limited contrast model was further shown by Tolhurst *et al* (1996) and Tolhurst and Tadmor (1997a,b). They demonstrated that the task of discriminating $\Delta\alpha$ did seem to rely upon contrast changes within a limited spatial frequency band by using stimuli in which the value of α had been changed only within restricted spatial frequency bands. However, it should be noted that these restricted bands were actually 2 octaves wide.

1.1 Alternative explanations to the local band-limited contrast model.

Despite the suggestive evidence, the local band-limited contrast model might not be the only possible explanation for discriminating changes in the slope of the amplitude spectra. Indeed, Rohaly *et al* (1997) have shown that many different models or variant models of discrimination sometimes produce very similar predictions. It is still arguable that the observer may somehow estimate the spectral slope for both the reference and the test stimuli, perhaps by comparing the energy or contrast at two or more different spatial frequency bands within each image. Figure 1A, B illustrates the two alternative explanations. In Fig.1A, it is presumed that the observer makes use of cues only in a narrow frequency band (as proposed by Tadmor and Tolhurst), ignoring the rest of the spectrum. In Fig.1B, it is presumed that observer makes use of the whole spectral range to estimate slope *per se*. A slope comparison might, for instance, result if the observer did

concentrate on the contrast in one band, but after contrast had been normalised by division by the contrast in some other frequency bands (Heeger 1992; Foley 1994; Rohaly *et al* 2000).

The purpose of this paper is to investigate the plausibility of the alternative model where observers might be estimating the changes in the amplitude slope by utilising Fourier energy at two or more spatial frequency bands. To do so we have designed an experiment where the observer's performance for a simple single-frequency-band contrast discrimination task should be severely disrupted. This is done by introducing different amounts of random variation into the overall contrasts of the reference and test images. Figure 1C,D shows the effect of this disruption on both models. As it can be seen in the Figure, the random shifting (upwards or downwards) of the amplitude spectra ought to vary the amounts of Fourier energy in a given spatial frequency band while leaving the amplitude slope the same. The disruptive effect of this contrast variation should be much greater in the case where the observer is estimating the spectral slope by discriminating changes in only one spatial frequency band.

2. Methods

2.1 Apparatus

The stimuli were displayed on a custom-made Joyce Electronics raster display (P4 phosphor) at a 100 Hz frame rate with a mean luminance of 45 cd.m^{-2} . The screen measured 30.3 cm by 22.9 cm, which was equivalent to 7.58 deg by 5.73 deg at the eye from the observer's viewing distance of 2.30 m from the display. The overall controlling system was an IBM compatible PC housing a VSG2/3 (Cambridge Research Systems) graphics card. This graphic card has a framestore memory that is coupled to *two* colour palettes, each with 8-bit digital-to-analogue converters (DAC), and 8-bit look-up tables (LUT). A simple 8-bit display with only 256 grey levels is not adequate for experiments in which it is necessary to change contrast in small steps while still retaining the integrity of

the picture or grating. For most of our experiments we needed subtle control of contrast. For these we used the Pelli & Zhang (1991) modification which allowed us to add together the two palette outputs to form a higher resolution output. This mode allows the choice of 256 grey levels on any one frame out of a total choice of 4096 grey levels (“pseudo 12 bit resolution”).

All images were viewed binocularly, in a darkened room with constant dim artificial lighting so that the display screen was viewed against a background with luminance of about 1 cd.m^{-2} . The observer was seated in a comfortable chair with a high neck support that ensured a fixed viewing distance of 2.3 m. Two image sizes were used, Size 1 (4cm x 4cm, or 1 deg x 1 deg) and Size 2 (8cm x 8cm, or 2 deg by 2 deg). Each stimulus was composed of 128 by 128 pixels whose size was about 0.5 and 1 minutes of arc on the retina for Sizes 1 and 2 respectively. Two observers participated; one was emmetropic, the other was a fully corrected myope.

2.2 Stimuli

Figure 2 shows the three photographs of natural scenes that were employed as stimuli, the same as in earlier work (Tolhurst & Tadmor, 1997b). A detailed account of how these were photographed and digitised can be found elsewhere (Tolhurst *et al* 1992, Tadmor and Tolhurst 1994). Once in a digital form, non-linearities of the negative’s optical density were corrected. Following this, the images were represented at 128 x 128 pixels and with 256 grey levels and were Fourier transformed. The resulting amplitude spectra were filtered so as to be consistent with Equation 1 with α values equal to 0.4 (whitened), 1.0 (near-normal) and 1.4 (pseudo-blurred). From each of these original images a set of stimuli was constructed. Each set consisted of a *reference* stimulus whose amplitude slope was equal to α and a sequence of 20 *test* stimuli whose amplitude slope was gradually increased ($\alpha + \Delta\alpha$) in steps of 0.02. The final images were obtained by performing an inverse Fourier transform on these altered spectra.

Modifying the slope of the amplitude spectrum in this way causes the overall power spectrum to change as well (Bracewell 1986). To avoid this spurious cue for discriminating between a reference and a test stimulus, all images within a set had their powers normalised to the same value before being inverse-Fourier transformed. This effectively means that,

when the negative slope (α) is increased, the line is simultaneously shifted upwards, so that the reduction in power at high spatial frequencies is compensated for by an increase in power at low spatial frequencies. This can be seen schematically in Fig.1A. Thus, this normalisation of the amplitude spectra differs from a true blurring operation, where the amplitude is decreased at all frequencies. The normalisation was carried out only within a given *set* of pictures, allowing different sets to have different powers. Figure 3 illustrates the differences between three different sets derived from the same photograph. Although some of the image quality is lost when printed on paper, one still can see the tangible differences produced by altering the α value. At $\alpha = 0.4$, high spatial frequencies are emphasised although the image appears rather shadowy and of low contrast. When α equals 1.0, the image appears closer to our everyday perception and is reasonably in focus. At $\alpha = 1.4$, the image looks softer and blurred, consistent with the lack of high frequencies.

2.3 Measurements of discrimination thresholds

2.3.1 Stimulus presentation.

To measure the thresholds for discriminating changes in α within a given *set* of images, we used a modified two-alternative forced-choice (2AFC) paradigm. A single trial consisted of presenting three images from the set successively for 500 ms, with 200 ms time intervals between them. Two of these three images were the *reference* image and were identical. The other was one of the *test* images from the set. The amplitude spectra for the two reference images was f^α while the test image had an amplitude spectra of $f^{(\alpha+\Delta\alpha)}$. The order in which the stimuli were presented was decided at random by the computer, except that the middle interval was always the reference. Since the middle image was always the reference, the observer's experimental task was to decide whether the test image (i.e. the odd one out) was presented in the first or the third interval. The onset of each image presentation was indicated by an auditory signal, and the observer's response was recorded by pressing the appropriate button (left or right) on a remote-control console. Auditory feedback was given as to whether the observer's choice was correct. The overall power of all three images was always the same and their mean luminance was the 45 cd.m⁻² the same as the immediately surrounding display. The stimulus was centred on the screen and the observer

was allowed to fixate freely. All other parts of the screen (not occupied by the stimulus) had a fixed luminance of 45 cd.m⁻².

2.3.2 Random contrast variation.

As described in the Introduction, the purpose of this work is to determine the plausibility of the local band-limited contrast model by attempting to disrupt the observer's performance for single frequency-band contrast discrimination. To do this, random contrast variation was imposed on the stimuli during presentation; a different amount of variation was imposed on the three stimulus presentations within a single trial. The amount of random contrast variation was specified as a percentage of the original contrast of the image ($\epsilon\%$). In practice, this means that the overall power (or r.m.s. contrast) of any image was reduced by an equi-probable random value between 0% and $\epsilon\%$. Consequently, all three images (including the two reference ones) presented on a single trial were now likely to differ from each other in their overall power.

2.3.3 Calculation of the discrimination threshold.

Once the set of images was chosen (e.g. one of the series in Fig.3), two different staircases were initialised. One was initialised with a *test* stimulus that was well below the discrimination threshold (i.e. a very low $\Delta\alpha$) and the other with a stimulus well above the discrimination threshold. Both staircases had the same *reference* stimulus. During a single experimental "run" each of the test stimuli were presented 5 times. If, for example, 5 out of 5 responses were correct, the value of $\Delta\alpha$ was decreased, making the task harder in the next trial. If the number of correct responses were 3 or fewer then $\Delta\alpha$ is increased, making the task easier. α was increased or decreased in linear steps. In general the two staircases converged within 10-15 individual runs (of 5 presentations for each staircase), and 20 runs were carried out. The discrimination threshold was estimated by fitting the measured psychometric function with the integral of a normal distribution, which was constrained to fall within 50% correct answers (the guess rate for this 2 AFC experiment) and 98% correct answers (allowing 2% of "finger error"). The discrimination threshold was defined as the value of $\Delta\alpha$ eliciting 74% correct responses and was found by interpolation on the fitted curve. The goodness of the fit was estimated by verifying that χ^2 was lower than the number of degrees of freedom in most of the cases. The standard error of the threshold and the value of the slope (β) of the fitted function were also estimated from the inverse of the second-differential of the merit function at the point where the parameters gave the

best fit (Edwards 1976). In fact, the thresholds for several experimental conditions were determined concurrently, by randomly interleaving the staircases for the different conditions.

2.4 Calculation of local contrast

We model the human observer's ability to discriminate changes in the amplitude spectra of complex scenes in terms of a relatively simple task: the discrimination of changes in the Michelson contrast of simple sinusoidal gratings. To do this we need to find an appropriate definition of *contrast*.

A physiologically-plausible definition should not only characterise contrast as a measure of local luminance modulation divided by the local mean luminance, but it should take into account how well the neurons or channels (which are optimally responsive to certain spatial frequencies) are activated by the stimulus. Consequently, this definition should be calculated at a series of different spatial scales (or spatial frequency bands) using operators whose properties broadly resemble those of visual neurons (see Peli, 1990; Tolhurst & Tadmor, 1997a,b).

We define local contrast at a given position in the picture $[x,y]$ and within a specified spatial frequency band F as:

$$C_F(x, y) = \frac{a_F(x, y)}{l_m(x, y)} \quad \text{Eqn.2}$$

where $a_F(x, y)$ is a *band-pass* filtered version of the image convolved with a circularly-symmetric operator which has the following spatial frequency characteristic:

$$A_F(f) = e^{-\left(\frac{(f-F)^2}{2\sigma^2}\right)} \quad \text{Eqn.3}$$

$l_m(x, y)$ is the result of *low-pass* filtering the image with a circularly-symmetric operator with the following spatial-frequency characteristics:

$$L_m(f) = e^{-\left(\frac{f^2}{2\sigma^2}\right)} \quad \text{Eqn.4}$$

f is the spatial frequency, F is the centre frequency of the particular contrast band and σ is the standard deviation of the Gaussian operator. This parameter is the same in Eqns 3 and 4, ensuring that both the “modulating” signal (a_F) and the mean-luminance signal (L_m) had the same spatial weighting and extent. This local-band contrast operator roughly represents an “on” centre ganglion cell. A_F represents the antagonistic centre-surround of a typically band-pass cell. The division by L_m represents local light adaptation (Shapley & Enroth-Cugell 1984). We set σ to be 0.3 times F , giving the operator a spatial frequency bandwidth at half height of about 1 octave (similar to that of human channels (Blakemore & Campbell 1969; Watson and Robson 1981) and neurons in primary visual cortex of a variety of mammalian species (Movshon *et al* 1978, Tolhurst & Thompson 1981, De Valois *et al* 1982, Baker *et al* 1998). The contrast in a frequency band at a single point in a picture is given by Eqn.2; the absolute value of contrast in the band was then averaged across the picture to give the “contrast in the band”.

2.5 Just noticeable differences in real contrast.

2.5.1 Sine-waves

To evaluate whether the observer’s discrimination thresholds for changes in α in complex stimuli can be explained as discrimination of changes in contrast for a given frequency band (F), we must first determine how well the observer is capable of detecting such changes for simple stimuli. This is done by measuring how well the observer can discriminate changes in the “real” Michelson contrast of simple sinusoidal gratings of frequency F and by evaluating whether the corresponding change in contrast in a similar frequency band (between the test and reference images) is large enough to be detectable.

The observer’s discrimination functions for sinusoidal gratings were not measured directly. We used a template corresponding to the familiar “dipper” function for contrast discrimination (Campbell and Kulikowski 1966, Nachmias and Sansbury 1974, Tolhurst and Barfield 1978, Legge and Foley 1980, Legge 1981; Foley, 1994). The position of this

template in the plot was determined by measuring the observer's contrast thresholds for *detecting* sinusoidal gratings of spatial frequency F . To achieve this, we used the same experimental procedure as before, since the detection threshold is equivalent to the discrimination threshold when the reference image has zero contrast.

Figure 4 shows how we obtained the predicted “dipper” function or contrast discrimination function for a given spatial frequency. Once the observer's contrast thresholds for *detecting* sinusoidal gratings of frequency F is determined (plot on the left), the predicted dipper function for gratings of the same spatial frequency is aligned. The low contrast asymptote and the lowest point of the dip are made equal to the detection threshold (right plot on Fig.4). All detection thresholds were measured using square patches of sinusoidal grating, 4x4 cm presented foveally for 500 ms and using a similar 2AFC paradigm to that described above. Michelson contrast of simple sinusoidal gratings was defined as:

$$Michelson_contrast = \frac{(L_{\max} - L_{\min})}{L_{\text{mean}}} \quad \text{Eqn.5}$$

Where L represents luminance. Once these measurements have been completed, we are in a position to determine whether the change in equivalent contrast within a certain spatial frequency band is actually large enough to be detectable by the observer.

2.5.2 Band-passed natural scenes

We tested the validity of this theoretical dipper function directly (Figs.10 and 11). As well as measuring the discrimination of contrast with sinusoidal gratings as the stimuli, we also used band-passed versions of the natural images (Fig.2) themselves. The digitised photographs were Fourier transformed and were filtered with an isotropic filter having a spatial-frequency bandwidth of about 1 octave. The test images were then constructed by inverse Fourier transformation. The contrast of the stimuli was calculated as in section 2.4 as the local contrast in the band centred on the filter used in the band-pass operation. We used the same 2AFC technique to measure the minimum contrast increment needed to detect a change in contrast of the band-passed images at a number of pedestal contrasts.

3. Results

3.1 The effects of contrast disruption on slope discrimination

As stated above, any change in spectral slope will be accompanied by a change in local contrast within certain spatial frequency bands. Our experiments were designed to disrupt the visual system's performance if the observer is detecting contrast changes in only one of these bands but to interfere very little if the observer is evaluating changes in the amplitude slope *per se* by comparison between bands.

Figure 5 shows the discrimination thresholds for changes in the amplitude slope ($\Delta\alpha$) when the random contrast disruption is applied. The ordinate represents range of the random contrast disruption ($\epsilon\%$) at each measurement (see Methods). The graphs show the results for two observers. The image numbers correspond to those of the images shown in Fig.2. The reference value of the slope α was 1.0 (close to natural statistics) in all cases. Observer JTCL showed no noticeable increment in the discrimination threshold even when a disruption of up to 100% of the contrast was added to the pictures (Pictures Size 2, left side plot). Observer CAP showed a small increment in the discrimination threshold (of about three times the average standard errors) when up to 75% contrast disruption was added (Pictures Size 1, right side plot).

Figure 6 shows similar results for reference α of 1.4. Both observers showed no noticeable (larger than the standard errors) increment in their discrimination thresholds despite massive contrast disruption. Again, observer JTCL (left plot) used pictures size 2 and observer CAP used pictures size 1. Thus, at α values of 1.0 (near normal images) and 1.4 (seemingly blurred images), the added contrast disruption did not cause a major disruption of the threshold for detecting changes in the amplitude slope ($\Delta\alpha$); this implies that the simple model of Tadmor & Tolhurst (1994) and Tolhurst & Tadmor (1997a,b) is *not* applicable in these situations.

However, Fig.7 shows the results for image series where the reference α was equal to 0.4 (edge enhanced or whitened images). In these series, the contrast disruption produced a remarkable effect in both observers for the two different image sizes. In the most extreme case (subject CAP, Image 3, size 1) thresholds were increased by a factor of about 5 times

(ten times the size of the standard error bar). These results imply that in the case of edge enhanced images, observers may be detecting changes in contrast in just one narrow spatial-frequency band. It is clear, at least, that a different strategy is employed for whitened images than for the series with steeper spectral slopes.

3.2 Contrast reduction vs. contrast disruption.

These results suggest that the band-limited local contrast model is not applicable in its simplest form and that, instead, some kind of slope analysis is being employed in this discrimination task. However, it is possible that these results are affected by the reduction in average contrast that necessarily accompanies the contrast disruption; the affects of the disruption itself may have been obscured. In order to dismiss this possible artefact, we performed a control experiment that follows the same experimental paradigm. In this case, instead of introducing a random contrast disruption we simply reduced the average contrast of our images by a determined *fixed* percentage; all three stimuli in a trial would now have the same (but low) contrast. Figure 8 shows the results for this control experiment. Both graphs show little change when contrast was reduced by a fixed amount, confirming that a mere reduction of the overall contrast has no effect on the observer's ability to discriminate changes in α .

3.3 Disruption of real contrast discrimination in gratings

Figures 5 and 6 showed that the addition of random contrast disruption disturbed the observer's performance very little at reference α values of 1.0 and 1.4, but was greatly disturbed at an α value of 0.4.. However, it is not immediately obvious how much threshold elevation would be expected on the scale of $\Delta\alpha$ in response to a disruption of contrast. For comparison and for "calibration" of the effect, we examined the effect of the same random kind of contrast disruption on the observer's ability to detect changes in real Michelson contrast of simple sinusoidal gratings. We performed experiments on gratings

with spatial frequency and contrast similar to those that we surmised that the observer had been using to make the discriminations in the complex images.

Table 1 lists the calculated contrasts in one limited (1-octave) spatial frequency band for the reference stimuli used in all of the above experiments. Three image sets (images 1, 3 and 10) and three reference values of α (0.4, 1.0 and 1.4) are considered. The frequency bands chosen as the most likely to have been involved in the discrimination task were 2.5 and 10 cycles/picture for observer CAP at different α values, and 4, 15 and 12.5 cycles/picture for observer JTCL. These spatial frequency bands were selected because the change in contrast in that band between the reference and threshold test images within a given image set lay closest to the observer's contrast discrimination function (as described in section 2.5).

Figure 9 shows two examples discrimination experiments using sinusoidal gratings of the spatial frequencies and pedestals shown in Table 1. The discrimination experiments were carried out using gratings of the same size as the images previously employed (size 2 for observer JTCL and size 1 for observer CAP). Both experiments show a very large increment in the contrast discrimination threshold for sinusoidal gratings when the disruption is applied. This shows the extent of the disruption expected when a visual discrimination is restricted to only one spatial frequency band.

To be able to compare the results with sinusoidal gratings (Fig.9) *directly* with those of the previous experiments on the discrimination of changes in the amplitude spectra of natural scenes, we must have a common threshold metric for the two kinds of experiment. It is conceivable that, say, a small change in $\Delta\alpha$ for a reference α of 1.4 corresponds to a large change in contrast. Table 2 shows how the discrimination thresholds ($\Delta\alpha$) for each image at each reference α value was increased when a contrast disruption of 75% was applied. It also shows how the calculated band-limited contrast difference (ΔC) between the (threshold) test and the reference stimulus varies. The spatial frequency chosen for analysis was the same as in Table 1. The observer considered here is CAP.

Table 2 shows that the increments of calculated band-limited contrast (ΔC) are actually bigger than the corresponding changes in $\Delta\alpha$. We can compare these calculated contrast increments with the elevation of the contrast threshold actually produced in the sinewave grating experiments. Table 3 shows the increments in contrast (ΔC) necessary to

discriminate between real reference and test gratings in the presence of up to 75% contrast disruption (e.g. Fig.9). As we can see, The contrast disruption causes a much bigger percentage change in threshold for simple gratings than it does for complex images with α values of 1.4 and 1.0 (Table 2). The elevation of threshold for detecting changes in spectral slope from α of 0.4 is as great as that found for simple gratings when both thresholds are expressed as change in contrast.

4. Discussion

Previous studies (Tadmor and Tolhurst, 1994; Tolhurst *et al* 1996; Tolhurst and Tadmor 1997a,b) seemed to provide evidence that an observer's ability to discriminate changes in the amplitude spectra of natural images could be modelled as a simpler task: i.e., that of discriminating changes in the Michelson contrast of simple sinusoidal gratings. For each observer and for each reference α (Eqn.1), it seemed that there was *one* spatial frequency band where the changes in contrast were almost large enough alone to explain the observer's performance. This conclusion was supported by experiments using stimuli in which the spectral slope was changed only within restricted frequency bands: the ability to perform the task was much more reliant on some frequency bands than on others (Tolhurst *et al* 1996; Tolhurst and Tadmor 1997a,b) However, it was pointed out that, given the spectral complexity of the stimuli, it would be surprising if the observer's discriminations should rely on only a limited portion of the available amplitude spectrum. In fact, the changes in band-limited contrast in complex stimuli were sometimes not large enough on their own to account for the magnitude of the discrimination threshold at a single spatial scale. Thus, there is scope for an alternative hypothesis. There might be some interaction or comparison between two or more spatial frequency bands.

In this paper, we report an experiment designed to determine whether this alternative view is correct or not. Our results suggest that the observer must somehow be detecting changes in the amplitude slope by using information from more than a single spatial frequency band. We applied random amounts of contrast or power variation to the natural

image stimuli and found that, for reference α values of 1.0 and 1.4, this disruption had little effect on discrimination. This surely means that the observer could *not* have been relying on contrast information in single frequency bands, since the same contrast disruption did cause the expected large disruption of contrast discriminations with simple sinusoidal gratings.

The logic of the experiment, however, neglects *contrast normalisation* (Heeger 1992; Foley 1994; Rohaly *et al* 1997). It is increasingly clear that single frequency channels or frequency bands cannot be considered as independent coding entities, especially when the stimuli contain energy in many different frequency bands (as here). Rohaly *et al* (1997) considered a number of single-channel and multiple-channel models for predicting the discriminability of objects in natural scenes, and they concluded that there was surprisingly little difference in the predictions of the different models, except that inclusion of contrast normalisation in any model made it much better. If contrast normalisation were absolute, it would completely negate the logic of our experiment: even though we might change the physical contrast of the image, the contrast normalisation would ensure that the "internal representation" of the stimulus did not actually change. However, the contrast normalisation cannot be so perfect: discrimination of whitened images (α of 0.4) and of grating contrast *was* disrupted by contrast randomisation. Furthermore, reduction in the physical contrast of all our stimuli was accompanied by a reduction in perceived contrast. It remains possible that contrast normalisation could be a means by which an observer might "compare" the contrast energy in two bands: normalisation simply means that contrast in one band is represented relative to contrast in more than one band.

The arguments of Tadmor and Tolhurst (1994) and Tolhurst and Tadmor (1997a,b) were based on the finding that the changes in calculated contrast within the test stimuli were similar to those needed to discriminate contrast for real sinusoidal gratings. We should consider whether the magnitudes of the discrimination thresholds in the present experiments are also compatible with those for discriminating changes in the contrast of simple sinusoidal gratings. We follow the procedures of the previous papers by investigating whether the change in equivalent contrast between the reference and test stimuli (calculated within a certain spatial frequency band) is actually large enough to be detectable by the observer. If this change is not large enough, it is unlikely that a simple contrast discrimination forms the basis of any discrimination task. Contrast discrimination with simple sinusoidal gratings is described by the characteristic "dipper" template (Fig.10);

the template is positioned on a graph where the ordinate represents just-detectable contrast difference and the abscissa represents reference or pedestal contrast, using the results of a sinusoidal grating detection experiment (as described in section 2.5.1). In Fig.10, we show the theoretical “dipper” for discriminating differences in contrast for gratings with a spatial frequency of 10 cycles/picture. The filled symbols show that the template *does* describe the contrast discrimination thresholds using gratings and also natural images band-pass filtered on a centre frequency of 10 cycles/picture. The open symbols show the changes in *calculated contrast* in an experiment with the natural image stimuli: i.e., the effects of contrast reduction on the slope discrimination thresholds with a reference α of 1.0. As the Figure shows, the results lie close to the dipper but they do not fit the dipper as closely as might be expected if the observer was performing a simple contrast discrimination and if the observer was relying solely on information within the frequency band that we have illustrated. We have performed this analysis at other frequency bands and the match of the calculated contrast differences to the dipper was generally worse than that illustrated in Fig.10. Thus, no single frequency band is adequate alone to account for the observer’s ability to detect changes in spectral slope from a reference α of 1.0. The observer must have used information from more than one frequency band, and the nature of the comparison must have made the observer’s ability resistant to our imposition of random contrast perturbations.

However, this is not the case for a reference α of 0.4 (whitened images). Figure 11 shows a similar plot for the contrast reduction experiment when the reference α is 0.4 and the modelled spatial frequency is 2.5 cycles/picture. As we can see from the open symbols in the Figure, most of the transformed data for the slope discrimination experiment with natural images *do* now lie very close to the dipper template. This is just the behaviour expected from an observer who is making the slope discrimination in the complex images as if discriminating the Michelson contrast of a sinusoidal grating of only the given spatial frequency. Thus, the observer might have been using just this single spatial frequency band to perform the task, and this may explain why the contrast disruption described in section 3.1 produced such a large effect in the observer’s performance when the reference stimuli had a spectral slope of 0.4.

5. Conclusions

The results shown here challenge the validity of the Tadmor & Tolhurst (1994a) band-limited contrast model for discrimination of natural images which differ only in their amplitude spectrum slope (α). We found that the observers' thresholds for this task generally remain unaltered in the presence of a random contrast component designed to severely impair an observer's ability to make contrast judgements in just one independent frequency band. Thus we conclude that a band-limited local contrast model can only describe the processes underlying the slope discrimination task if there is some form of slope analysis involved. This might involve a comparison of local contrast between two or more spatial frequency bands. However, one set of results (whitened images) suggests that the observers may, indeed, be relying on only one spatial frequency band to perform the slope discrimination task.

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Table 1:

Reference α	Observer: CAP			Observer: JTCL		
	Frequency band c/picture	Picture no.	Band-limited contrast	Frequency band c/picture	Picture no.	Band- limited contrast
0.4	2.5	1	0.0214	4	1	0.0233
		3	0.0108		3	0.0143
		10	0.0132			
1.0	10	1	0.1016	15	1	0.1014
		3	0.0564		3	0.0537
		10	0.1242			
1.4	10	1	0.0880	12.5	1	0.0829
		3	0.0849		3	0.0708
		10	0.1199			

Table 1: The band-limited contrasts of the reference stimuli (3 image sets at three reference alphas) are shown at the 1-octave spatial frequency band most likely to have been involved in the discrimination, for each of the observers.

Table 2:

stimuli	Frequency (c/image)		Thresholds image.1		Thresholds image.3		Thresholds image.10	
			$\Delta\alpha$	ΔC	$\Delta\alpha$	ΔC	$\Delta\alpha$	ΔC
$\alpha = 0.4$	2.5	0% disruption	0.134	0.0081	0.074	0.0022	0.091	0.0032
		75% disruption	0.206	0.0136	0.348	0.0135	0.317	0.0147
		<i>threshold elevation</i>	53.7%	67.9%	370%	514%	248%	359%
$\alpha = 1.0$	10	0% disruption	0.121	0.0083	0.123	0.0037	0.101	0.0080
		75% disruption	0.158	0.0111	0.232	0.0092	0.187	0.0178
		<i>threshold elevation</i>	30.6%	33.7%	88.6%	149%	85.1%	123%
$\alpha = 1.4$	10	0% disruption	0.086	0.0118	0.102	0.0124	0.090	0.0160
		75% disruption	0.093	0.0131	0.160	0.0195	0.125	0.0219
		<i>threshold elevation</i>	8.14%	11.0%	56.9%	57.3%	38.9%	36.9%

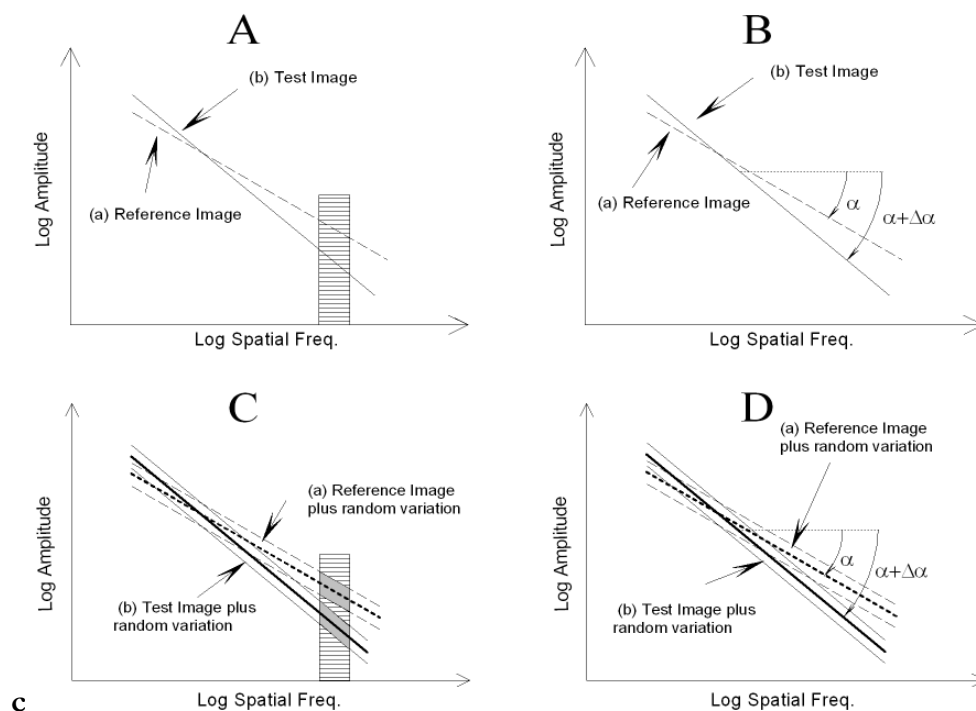
Table 2: The discrimination thresholds ($\Delta\alpha$) are shown for each of 3 image sets at each of 3 reference alphas, with and without an amplitude disruption of 75%. The band-limited contrast differences between test and reference stimulus at threshold are also shown, calculated for the spatial frequency band stated in column 2. The percentage elevation of threshold is expressed in terms both of slope (α) and of band-limited contrast (C).
Observer: CAP.

Table 3:

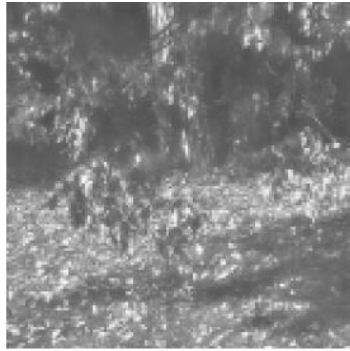
contrast	frequency (c/picture)		Threshold	
			$\Delta\alpha$	ΔC
0.031	2.5	0% disruption	---	0.0089
		75% disruption	---	0.0257
		threshold elevation		189%
0.056	10	0% disruption	---	0.0169
		75% disruption	---	0.0825
		threshold elevation		388%

Table 3: Elevation of the thresholds for discriminating differences in contrast between sinusoidal gratings of the spatial frequency and pedestal contrast shown. Observer CAP.

8. Figures

**Figure 1:**

The local band-limited contrast model and its alternative explanation. **A,B** The amplitude spectrum of a reference picture and a test picture are shown schematically on double logarithmic axes. Note that the two spectral lines cross at an intermediate frequency; power normalisation was carried out to ensure that the observer could not analyse changes in power instead of changes in amplitude slope (α). **A**, in the model of Tadmor & Tolhurst, the observer is presumed to evaluate the Fourier energy or contrast within a single spatial frequency band (hatched area). **B**, the alternative model consists in estimating the slope (by comparing the energy at two or more frequency bands). **C,D**, Effect of random perturbations in the overall contrast in both models. The effect of this variation should be greater on the observer's performance if he/she is evaluating changes between the stimuli by discriminating changes in only one frequency band (**C**). If the observer really does detect slope, then overall contrast perturbations will have no effect on performance, since the perturbation does not change the spectral slopes, only their positions on the y-axis (**D**).



Picture 1



Picture 3



Picture 10

Figure 2:

The original digitised photographs which were used as the basis of the stimuli used in this work. From left to right, Picture 1 (“tree”), Picture 3 (“portrait”) and Picture 10 (“street”).

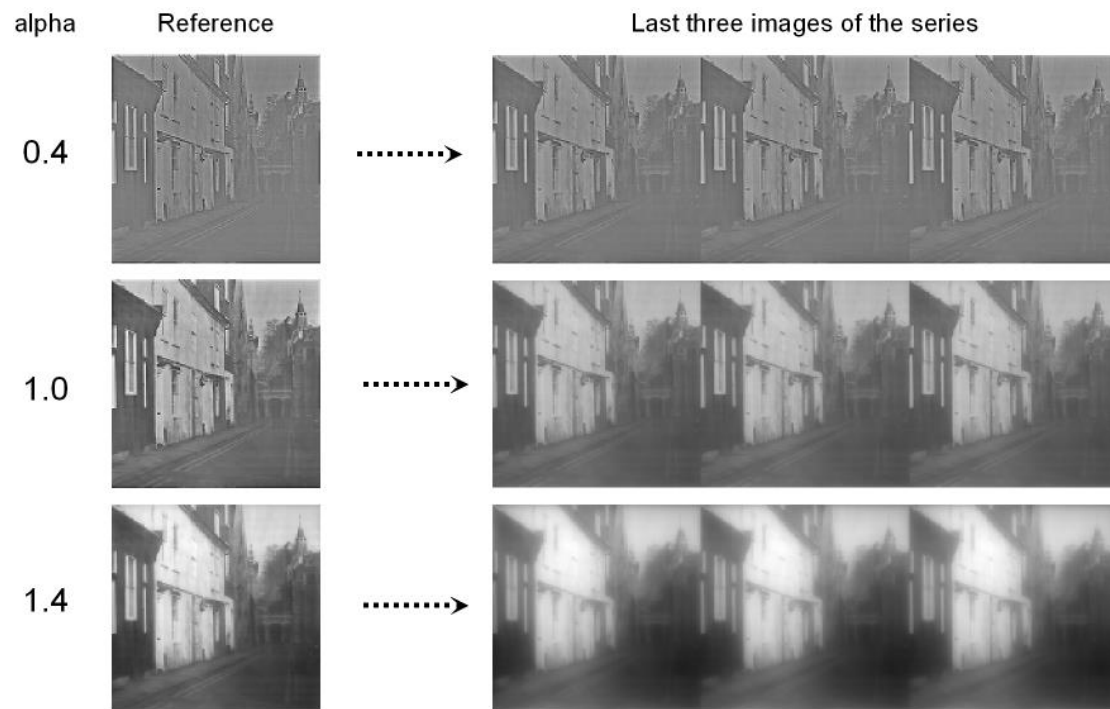


Figure 3:

Examples of stimuli made from the three different sets of images based on the same original photograph (picture 10). The value of α for the reference in the three rows is equal to 0.4, 1.0 and 1.4. The Figure shows the reference and the three last images of each set.

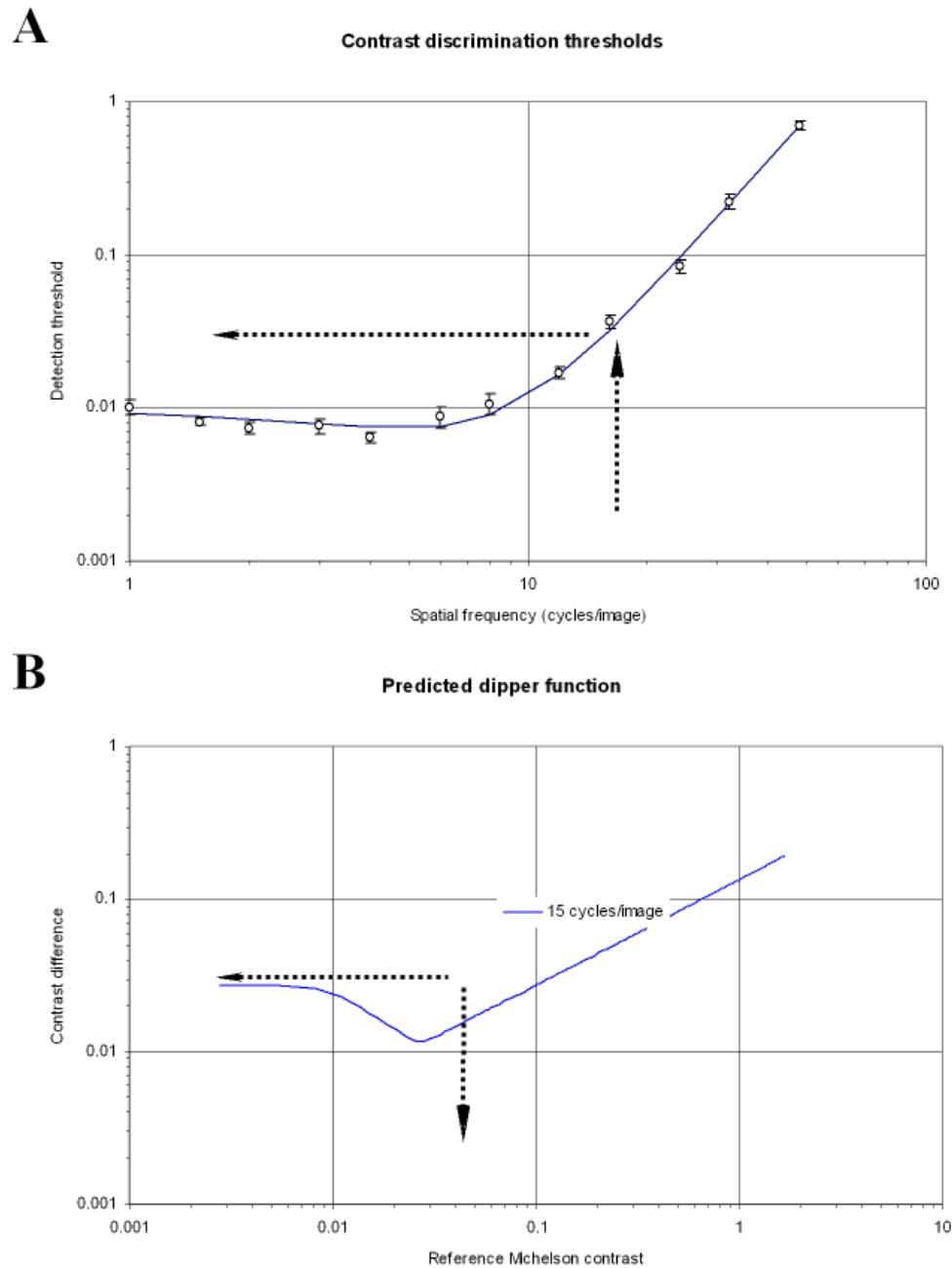


Figure 4:

A, the contrast thresholds for detecting a sinusoidal grating as a function of spatial frequency. In this plot the threshold corresponding to 15 cycles/image has been marked with an arrow. The second plot (**B**) shows how this threshold was used to position the predicted “dipper” for that spatial frequency on a graph of contrast discrimination threshold plotted against reference contrast; the arrows on abscissa and ordinate point at the contrast threshold from **A**.

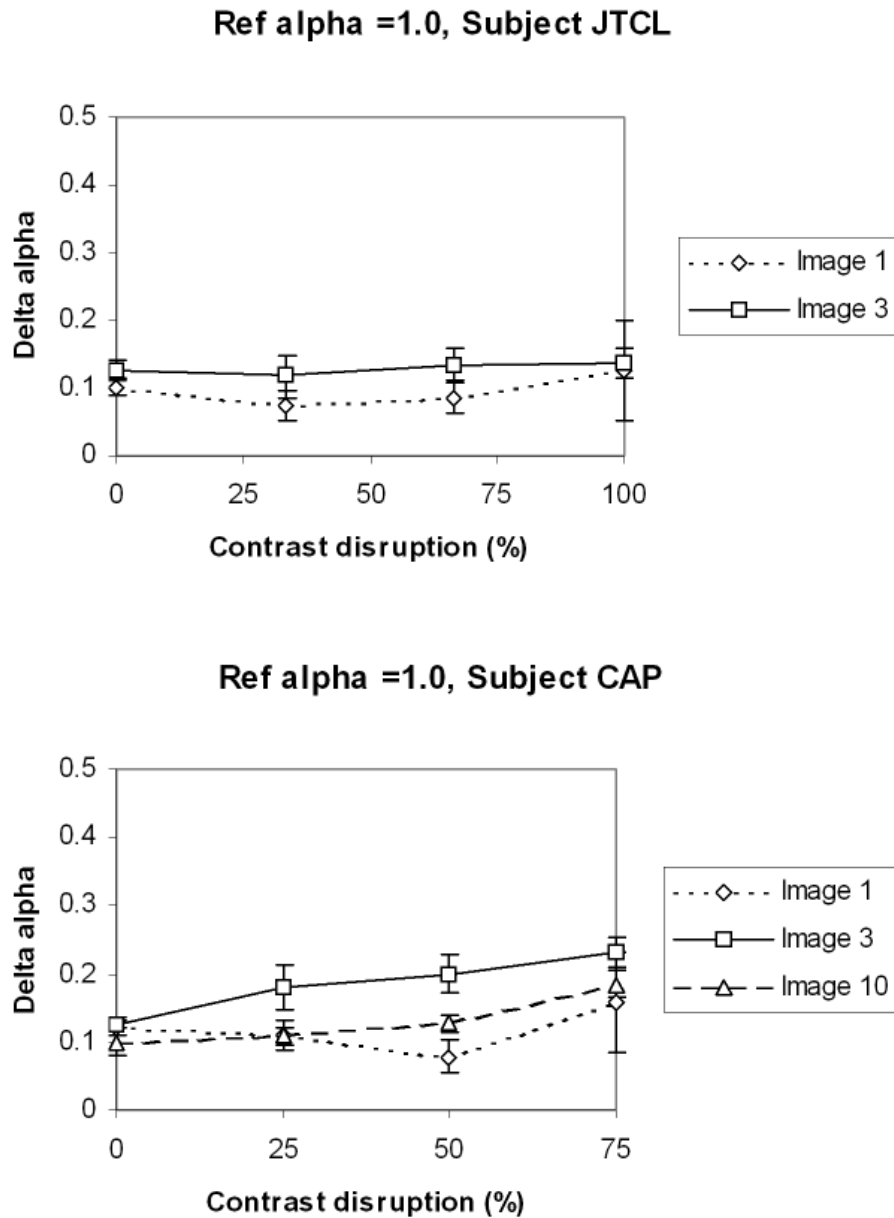
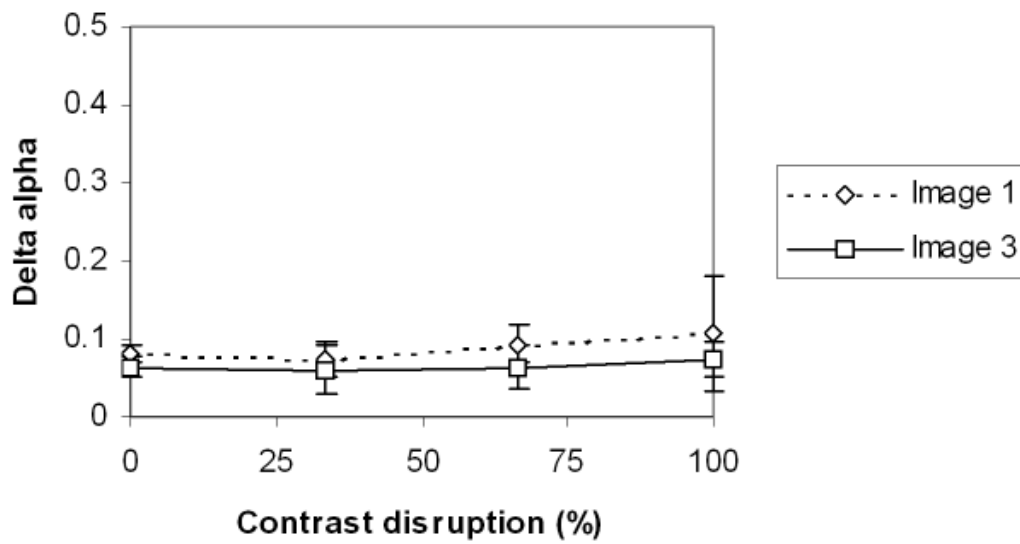


Figure 5:

Influence of different degrees of contrast disruption on the discrimination thresholds ($\Delta\alpha$). Two observers participated in the measurements: **A** - JTCL, **B** - CAP. Image sequences from all 3 original images were tested on both observers; JTCL was tested with size 2, CAP with size 1. The results here and in the rest of the paper are much the same for the two observers, despite the difference in image size. Reference α values were equal to 1.0. Image 1 – diamonds; image 3 – squares; image 10 – triangles. Error bars are +/- one standard error.

Ref alpha =1.4, Subject JTCL



Ref alpha =1.4, Subject CAP

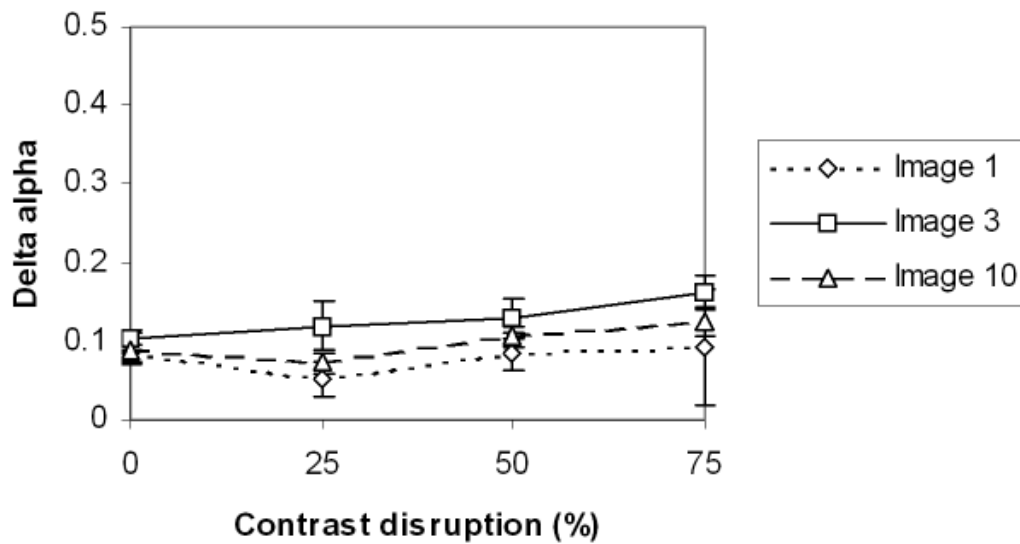


Figure 6:

The same as Fig.5, but for a reference α value of 1.4. **A** - JTCL, **B** - CAP.

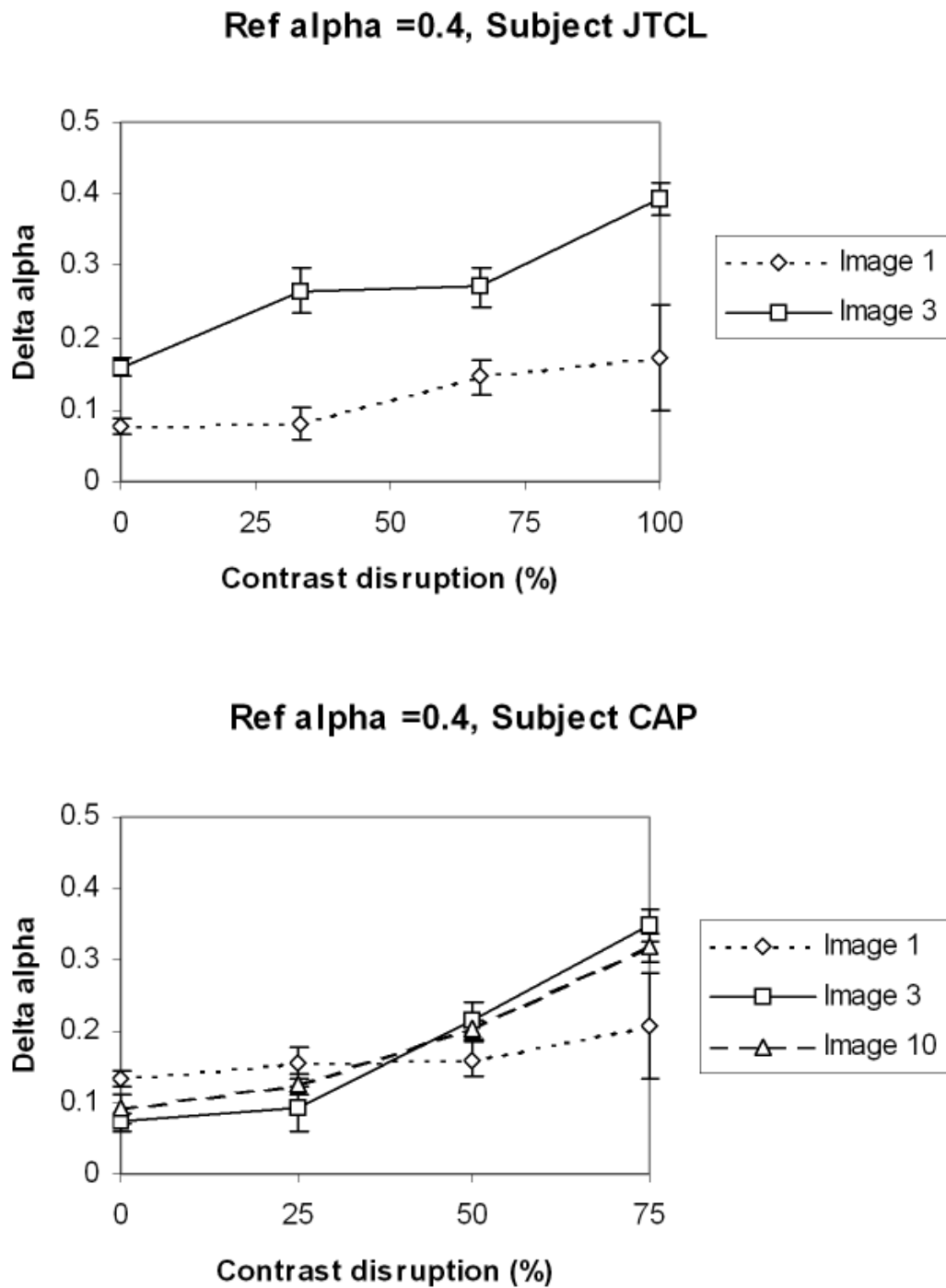
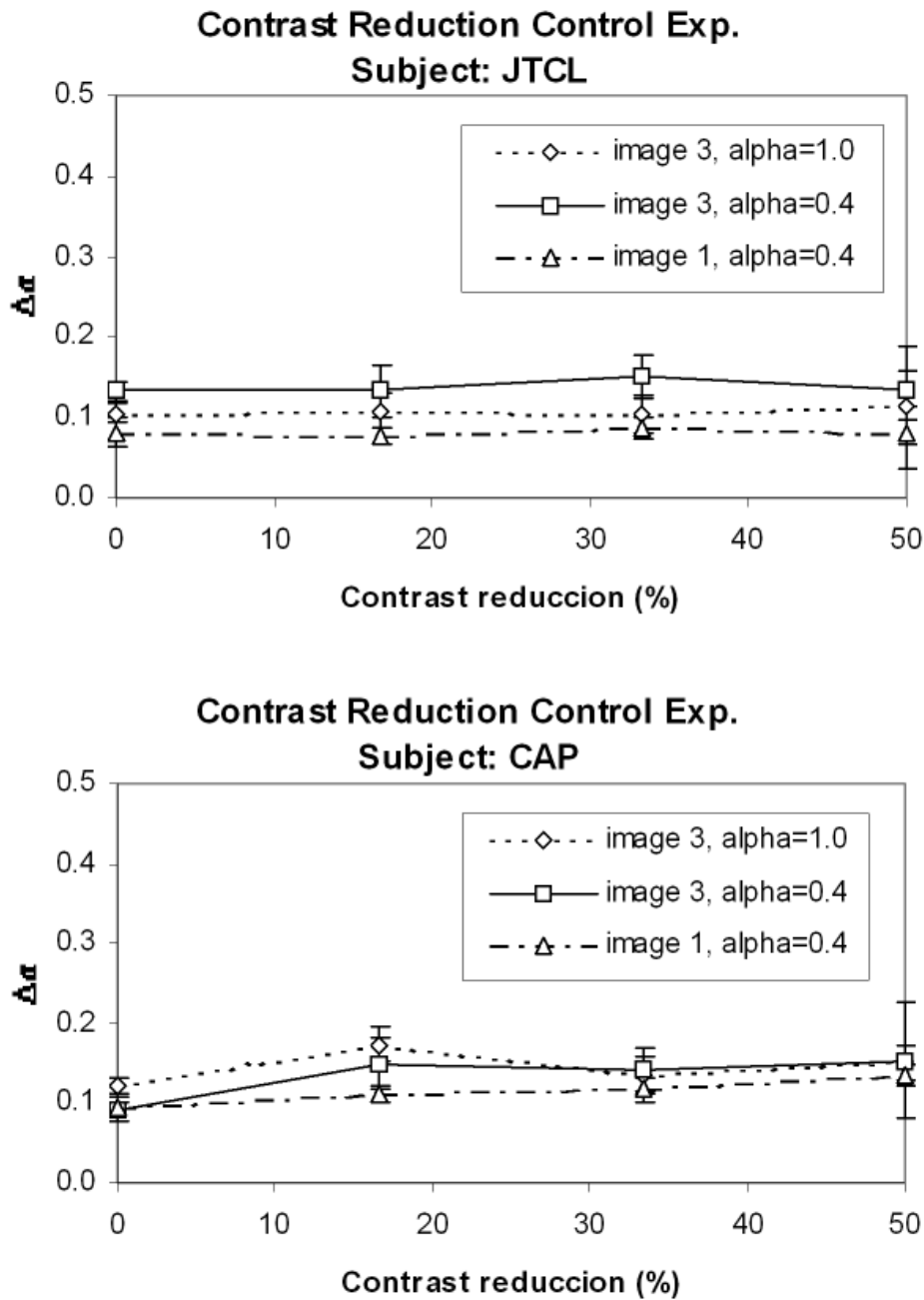


Figure 7:

The same as Fig.5 but for a reference α value of 0.4. Note that the results with these whitened images have a different character to those in Figs.5 and 6. **A** - JTCL, **B** - CAP.

**Figure 8:**

Results for a control experiment showing that reducing the contrast by a fixed amount (but without random disruption) does not produce a significant change in the observer's ability to discriminate variations in α . **A**, Subject JTCL used images size 2 and **B**, subject CAP used images size 1. Results are shown for picture 1, reference $\alpha = 0.4$ (triangles); picture 3, $\alpha = 0.4$ (squares); and picture 3, $\alpha = 1.0$ (diamonds).

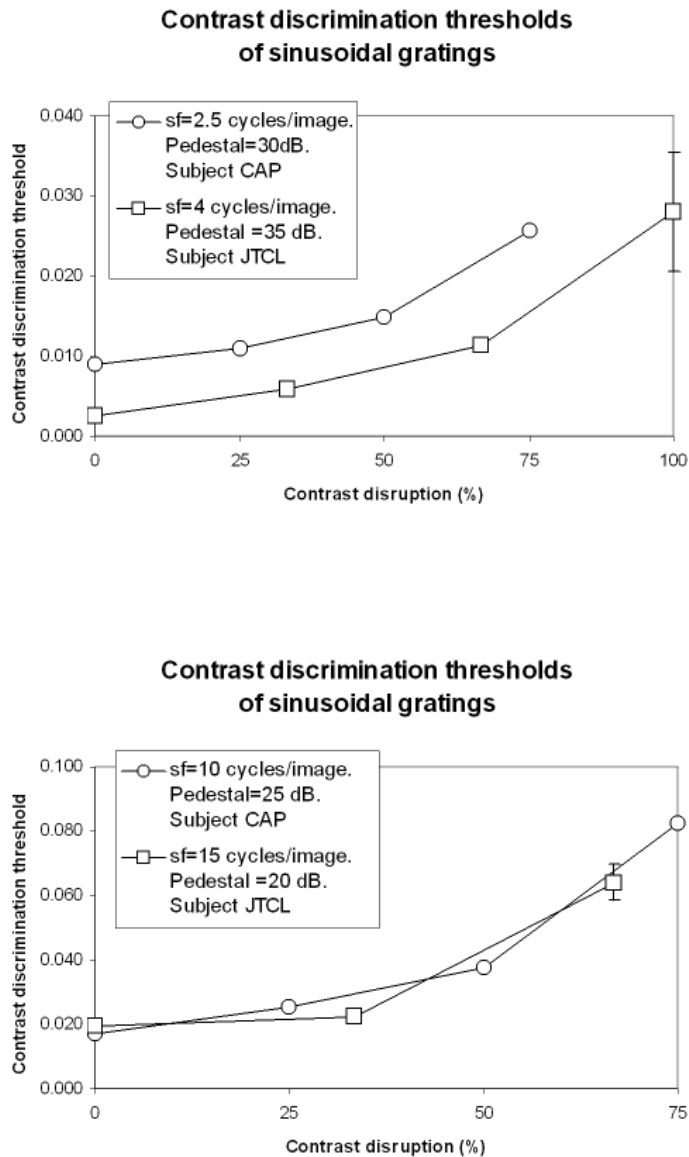


Figure 9:

The effect of contrast disruption on contrast discrimination thresholds for sinusoidal gratings. **A**, the results for gratings of 2.5 cycles/picture (observer CAP, circles) and a pedestal contrast of 0.032, and 5 cycles/picture (observer JTCL, squares) and a pedestal contrast of 0.018. **B**, the results for gratings of 10 cycles/ picture (observer CAP, circles) and a pedestal contrast of 0.056, and 15 cycles/picture (observer JTCL, squares) and a pedestal contrast of 0.1.

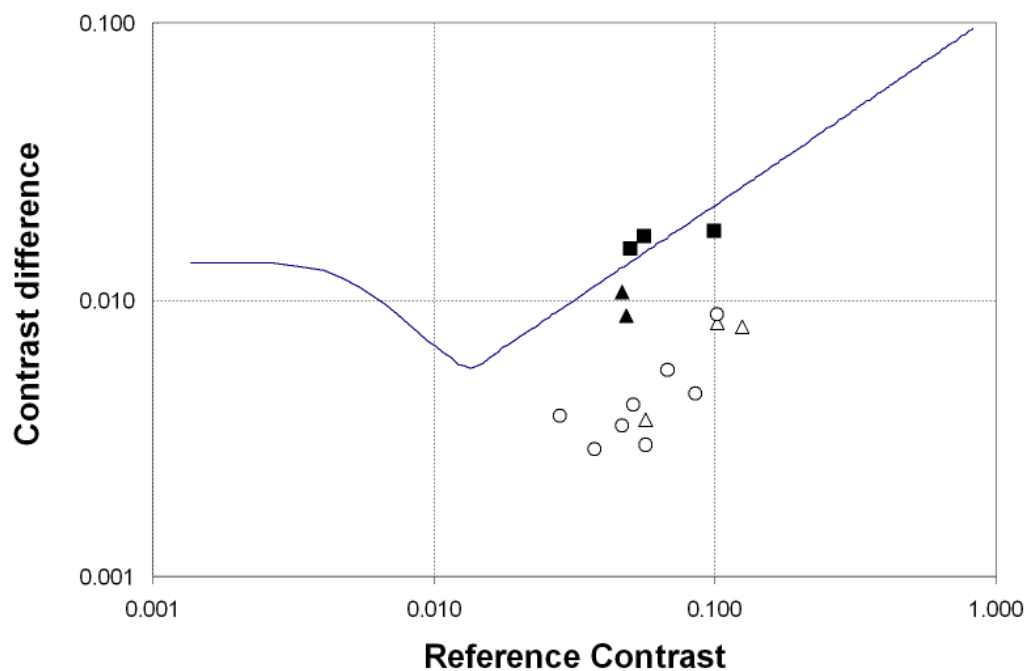


Figure 10:

Modelling slope discrimination at reference $\alpha = 1.0$ by contrast discrimination at 10 cycles/picture. The solid line represents the theoretical “dipper” function for contrast discrimination for simple sinusoidal gratings with spatial frequency of 10 cycles/pic, given the observer’s measured contrast detection threshold for a grating of this frequency. The filled squares show actual measurements of the just-noticeable difference in contrast for real gratings of 10 cycles/pic. The filled triangles represent jnds for contrast discrimination using band-passed images (pictures 1 and 10); the pass band was about 1 octave wide, centred on 10 cycles/pic. The open circles and open triangles show the *calculated* contrasts in the reference images and the *calculated* contrast differences at threshold in the experiment where the observer was obliged to detect differences in spectral amplitude slope from a reference α of 1.0. Contrast is calculated in the band centred on 10 cycles/pic, and the results modelled are from the contrast reduction experiment (images 1, 3 and 10 have not been distinguished in the plots to avoid unnecessary complexity). Observer CAP.

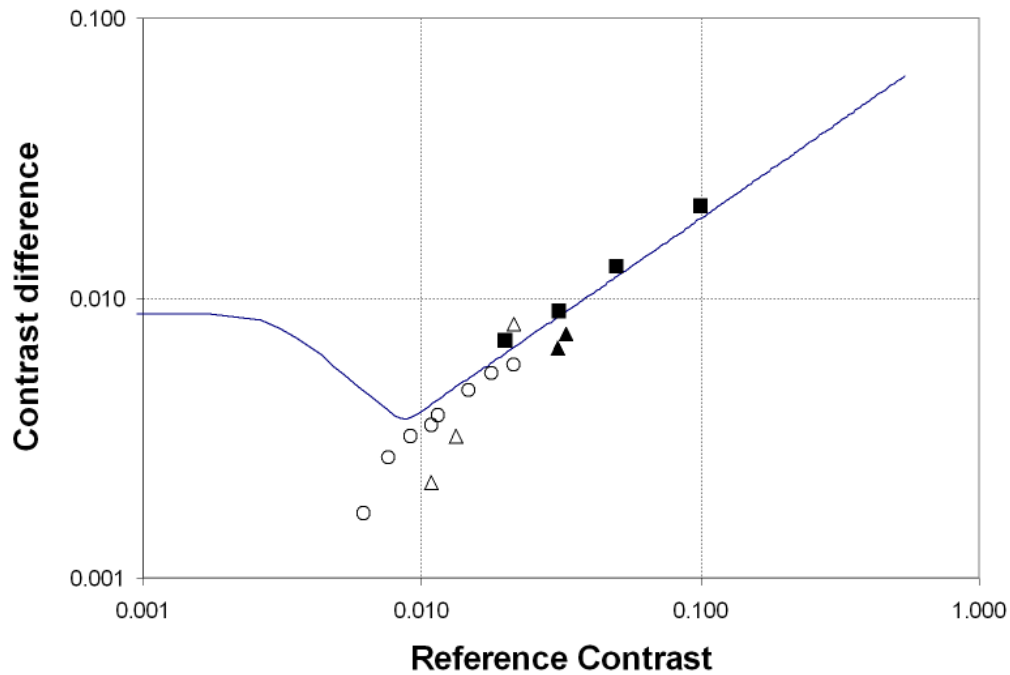


Figure 11:

As Fig.10, but modelling slope discrimination at reference $\alpha = 0.4$ by contrast discrimination at 2.5 cycles/pic. The solid line represents the theoretical dipper function for contrast discrimination of simple sinusoidal gratings with spatial frequency of 2.5 cycles/pic. The filled squares are the jnd results for real gratings of 2.5 cycles/pic, and the filled triangles represent jnds for images 1 and 10 bandpassed on a centre frequency of 2.5 cycles/pic. The open circles and open triangles show the *calculated* contrasts in the reference images and the *calculated* contrast differences at threshold in the experiment where the observer was obliged to detect difference in spectral amplitude slope from a reference α of 0.4. Contrast is calculated in the band centred on 2.5 cycles/pic, and the results modelled are from the contrast reduction experiment (images 1, 3 and 10 again are not differentiated in the plots). Observer CAP.