# Biomass production efficiency controlled by management in temperate and boreal

- 2 ecosystems
- 3 M Campioli 1\*, S Vicca 1, S Luyssaert 2, J Bilcke 3, E Ceschia 4, FS Chapin III 5, P Ciais 2,
- 4 M Fernández-Martínez 6-7, Y Malhi 8, M Obersteiner 9, D Olefeldt 10, D Papale 11, SL Piao
- 5 12-13, J Peñuelas 6-7, PF Sullivan 14, X Wang 12, T Zenone 15-16, IA Janssens 1

6

- 7 1 Centre of Excellence PLECO (Plant and Vegetation Ecology), Department of Biology,
- 8 University of Antwerp, 2610 Wilrijk, Belgium
- 9 2 LSCE CEA-CNRS-UVSQ, Orme des Merisiers, F-91191 Gif-sur-Yvette, France
- 10 3 Centre for Health Economics Research and Modeling Infectious Diseases, Vaccine and
- 11 Infectious Disease Institute, University of Antwerp, Antwerp, Belgium
- 4 Centre d'Etudes Spatiales de la BIOsphère, 31401 Toulouse cedex 9, France
- 13 5 Institute of Arctic Biology, University of Alaska Fairbanks, Fairbanks, AK 99775, USA
- 14 6 CSIC, Global Ecology Unit CREAF-CEAB-CSIC-UAB, Cerdanyola del Vallès, 08193
- 15 Barcelona, Catalonia, Spain
- 16 7 CREAF, Cerdanyola del Vallès, 08193 Barcelona, Catalonia, Spain
- 17 8 Environmental Change Institute, School of Geography and the Environment, University of
- 18 Oxford, Oxford OX1 3QY, UK
- 19 9 International Institute for Applied Systems Analysis, Schlossplatz 1, 2361 Laxenburg,
- 20 Austria
- 21 10 Department of Renewable Resources, University of Alberta, Edmonton, AB T6G 2H1,
- 22 Canada
- 23 11 DIBAF, University of Tuscia, 01100 Viterbo, Italy
- 24 12 College of Urban and Environmental Sciences, Peking University, Beijing 100871, China

This is the accepted version of the following article: Campioli, M. et al. "Biomass production efficiency controlled by management in temperate and boreal ecosystems" in Nature geoscience, vol. 8, no. 11 (Nov. 2015), p. 843-6, which has been published in final form at DOI 10.1038/ngeo2553

- 25 13 Institute of Tibetan Plateau Research, Chinese Academy of Sciences, Beijing 100085,
- 26 China

- 27 14 Environment and Natural Resources Institute, University of Alaska Anchorage, Anchorage,
- 28 AK 99508, USA
- 29 15 Department of Environmental Sciences, University of Toledo, Toledo, OH 43606, USA
- 30 16 present address: Centre of Excellence PLECO (Plant and Vegetation Ecology),
- 31 Department of Biology, University of Antwerp, 2610 Wilrijk, Belgium.
- 32 \* corresponding author

# **Summary paragraph**

34

35

36

37

38

39

40

41

42

43

44

45

46

47

48

49

50

51

Plants acquire carbon through photosynthesis to sustain biomass production, autotrophic respiration, and production of non-structural compounds for multiple purposes<sup>1</sup>. The fraction of photosynthetic production used for biomass production, the biomass production efficiency<sup>2</sup>, is a key determinant of the conversion of solar energy to biomass. In forest ecosystems, biomass production efficiency was suggested to be related to site fertility<sup>2</sup>. Here we present a global database of biomass production efficiency from 131 sites compiled from individual studies using harvest, biometric, eddy covariance, or process-based model estimates of production - dominated, however, by data from Europe and North America. We show that instead of site fertility, ecosystem management is the key factor that controls biomass production efficiency in terrestrial ecosystems. In addition, in natural forests, grasslands, tundra, boreal peatlands and marshes biomass production efficiency is independent of vegetation, environmental and climatic drivers. This similarity of biomass production efficiency across natural ecosystem types suggests that the ratio of biomass production to gross primary productivity is constant across natural ecosystems. We suggest that plant adaptation results in similar growth efficiency in high and low fertility natural systems, but that nutrient influxes under managed conditions favour a shift to carbon investment from the belowground flux of non-structural compounds to aboveground biomass.

52

53

54

55

56

57

58

# Main text

The fraction of gross primary production (GPP) used for biomass production (BP) of terrestrial ecosystems has recently been coined biomass production efficiency (BPE)<sup>2</sup>. BPE is typically used as a proxy for the carbon-use efficiency or NPP-to-GPP ratio, where NPP refers to net primary production i.e. BP plus the production of non-structural organic compounds<sup>1</sup>.

Current knowledge about BPE is mainly derived from research on forests. Earlier work reported BPE to be conservative across forests<sup>3</sup>, whereas more recent syntheses suggest high inter-site variability<sup>2,4</sup>. The variation in BPE was first attributed to vegetation properties (forest age) and climate variables<sup>4</sup>. More recently, it was shown that forest BPE in a range of natural and managed sites was correlated with site fertility, with management as a secondary BPE driver<sup>2</sup>.

Fertility and management are strongly correlated as management enhances productivity by increasing plant-available resources, including nutrients. For instance, fertilization of grasslands directly increases the ecosystem nutrient stock, whereas forest thinning indirectly increases nutrient availability at the tree level by reducing plant-plant competition. In addition, fertile sites are more likely than infertile sites to be managed. Atmospheric deposition of nutrients, especially nitrogen (N), might further complicate the relationship between BPE, fertility and management. The influence of site fertility and management on BPE has not been disentangled in previous studies, and the impact of N deposition on BPE is largely overlooked. Here, we postulate that the impact of management on BPE is underestimated. In addition to a direct effect on BPE through selection of the most efficient plants<sup>2,5</sup>, management can indirectly affect BPE through effects on site fertility and related belowground dynamics<sup>2</sup>. Understanding of these dynamics not only will clarify the controls of BPE but also elucidate the human impacts on BPE.

We compiled a new BPE dataset comprising 131 sites, including forests, grasslands, croplands, wetlands (temperate marshes and boreal peatlands) and tundras (Methods). All major climatic zones (from polar to tropical) were represented but managed sites were located almost entirely in the temperate and boreal zone of North America and Europe (Supplementary Fig. 1, Supplementary Table 1). For each site, our dataset also included vegetation characteristics, environmental data and information on anthropogenic impacts such

as management and atmospheric N deposition (Supplementary Table 2). With regard to management, we adopted a binary classification (Methods), distinguishing natural sites (pristine sites or sites with a low human impact that largely reproduced naturally occurring processes, e.g. grasslands with low grazing) from managed sites (sites dominated by human activity with impacts that would not occur in nature, e.g. newly established and fertilized grasslands). The utility of this classification was tested against more complex classifications (Methods), whereas its reproducibility was assured by the definition of several sub-categories within the 'managed' and 'natural' classes (Supplementary Table 3). The BPE dataset, comprising the ancillary site information, is available in Supplementary Data. Our data analysis consisted of (i) multinomial ordered logistic regressions to examine the relationship between fertility and management (code available in Supplementary Information) and (ii) linear (univariate analysis, multiple linear regressions) and non-linear approaches (Random Forest) to extract emerging relationships between BPE and its potential predictors (Methods).

The analysis proceeded in five steps, using different sub-sets of our database. (1) We analyzed all natural sites to test whether BPE is driven by natural variation in site fertility. The results showed that this hypothesis was not true. First, BPE did not differ significantly (p=0.83) among natural ecosystem types of contrasting fertility status i.e. tundra and boreal peatlands (nutrient-poor), temperate marshes (nutrient-rich) and forests and grasslands (with variable but overall intermediate fertility status), showing an average BPE (and s.e.m) of 0.46±0.01 (Figure 1; Supplementary Table 4). Second, the impact of fertility on the BPE of natural ecosystems remained non-significant when accounting for variation in fertility among forests (p=0.24, n=43), grasslands (p=0.72, n=16) or all natural sites lumped together (p=0.23, n=75; Supplementary Fig. 2). (2) We analyzed the relationship between fertility and management in natural and managed forests to verify their correlation and disentangle (i) the impact of management on fertility from (ii) the fertility status unrelated to management. This

analysis confirmed that management was a significant explanatory variable for site fertility (likelihood ratio test of models with and without management as covariate: chi-square=17.33, p=0.00017), whereas the relationship between N deposition and fertility was weak (likelihood ratio test: chi-square=4.80, p=0.091). This led us to model fertility as a function of management (taking into account that the fertility status was the result of both the impact of management operations on soil nutrient availability and the management choice of which land, e.g. high or low fertility, to manage) and to obtain model residuals for each site representing the 'fertility status not explained by management' and defined hereafter as 'unexplained natural fertility' (Methods). (3) Once the effect of fertility and management were disentangled, we evaluated their relative importance as controllers of BPE and compared them to other possible BPE drivers (e.g. vegetation and environmental characteristics, N deposition) within the forest dataset. This analysis revealed that management was the key determinant of the differences in BPE among forests, N deposition was the second most important driver, and the unexplained natural fertility was insignificant (Supplementary Table 5, Supplementary Fig. 3). The analysis also showed that stand age had a significant (negative) impact on BPE which however became negligible when compared to the effect of management and N deposition (Supplementary Table 6). (4) We compared the BPE of key natural and managed ecosystem types (grasslands, forests and croplands) that typically share similar environmental characteristics and are regularly converted into one another, and observed that the BPE of managed sites was substantially greater than the BPE of natural sites (Figure 1, Figure 2; Supplementary Table 7). (5) Last, we studied the impact of the potential drivers of BPE on all natural ecosystems and found that BPE of natural unmanaged sites was independent not only of the observed site fertility (see above point 1) but also of N deposition and largely independent of all the vegetation and environmental drivers examined (Supplementary Table 8, Supplementary Fig. 3). Climate showed an influence on BPE but this

109

110

111

112

113

114

115

116

117

118

119

120

121

122

123

124

125

126

127

128

129

130

131

132

effect was weak (0.05<p<0.10) and not consistent across statistical methods (Supplementary Table 8, Supplementary Fig. 3).

134

135

136

137

138

139

140

141

142

143

144

145

146

147

148

149

150

151

152

153

154

155

156

157

158

The observed positive impact of management on BPE does not come as a surprise in itself. Rather, the novelty of this study is the finding that management is by far the 'key' driver of BPE and more important than any other vegetation or environmental factors. This observation calls for a refinement of the hypothesis, which previously postulated that greater BPE in more fertile sites is related to reduced C allocation to symbiotic fungi, as plants in nutrient-rich conditions invest less in processes facilitating nutrient uptake<sup>2</sup>. Our revised hypothesis relies on the fact that adaptation processes in natural ecosystems<sup>6</sup> could allow plants in both nutrient-poor and nutrient-rich environments to have similar growth efficiency. However, belowground C transfers to symbionts are not static<sup>7</sup> and the greater nutrient availability caused by management could make root symbiotic associations less important for plants and thus reduce the flux of C from plants to symbionts. This pattern would favor C investment in biomass production, particularly aboveground, since light may become the most limiting resource. This hypothesis is supported by (i) the allocation pattern available for a subset of our forests showing that management substantially increased allocation to aboveground wood BP (+13%, p<0.001) and marginally decreased allocation to fine root BP (-4%, p=0.083) (Table 1) and by (ii) forest C allocation meta-analyses<sup>8</sup> which reported increased C partitioning to above ground BP and decreased partitioning to below ground C flux in response to fertilization. Declines in mycorrhizal fungi following fertilization are well known<sup>9</sup>. Similarly, thinning can negatively affect the standing crop of mycorrhizal fungi<sup>10</sup> and ectomycorrhizal metabolic activity<sup>11,12</sup>, which is consistent with our new interpretation. In addition, the larger BPE in managed ecosystems might also reflect decreased allocation of GPP to autotrophic respiration (Ra), thus lower Ra-to-GPP ratio<sup>2</sup>. However, as previous research does not support this hypothesis<sup>3,8</sup> and the variability of the Ra-to-GPP ratio might be

small, ad-hoc experiments combining the assessment of C transfer to mycorrhizal fungi and ecosystem Ra will be needed to ascertain the importance of these dynamics in managed ecosystems. Similarly, further research should explore (i) if the hypothesized reduction in C allocation to mycorrhizae (and exudates) might have a long-term negative feedback on the site nutrient availability where management does not include external input of nutrients and (ii) the impact of ecosystem degradation on BPE, especially in tropical areas that are often overexploited.

Nitrogen deposition also appeared to have a positive effect on BPE. Like management, elevated N deposition represents an artificial change in natural fertility and a perturbation of the nutrient cycle. The apparently contrasting evidence that N deposition does not affect BPE of natural ecosystems (when considered separately from the managed ecosystems) is likely related to the intensity of the deposition and the fact that N deposition might influence BPE (like other ecosystem processes<sup>13</sup>) only at higher deposition rates. Natural sites are typically found in less urbanized locations and in our dataset they were characterized by deposition rates 43% lower than those of managed ecosystems. Furthermore, adaptation responses to N deposition are more likely to occur in natural ecosystems where succession is much longer than rotations in managed ecosystems.

Little information was previously available about BPE of non-forest ecosystems<sup>14</sup>. Our analysis showed that BPE of natural ecosystems is independent of ecosystem type, vegetation and environmental characteristics (including natural site fertility). The lack of sensitivity of BPE to these potential drivers points to a rather conservative BPE across natural ecosystems. Our study supports the (highly debated) physiological argumentation for a constant ratio between BP and GPP in natural ecosystems<sup>3,4</sup> and provides important constraints for the global models that simulate high variability in BPE or NPP-to-GPP ratio.

Finally, our findings have practical applications, particularly for Europe and North America. (i) The quantification of BPE for managed ecosystems can improve yield simulations by models (e.g. timber in forests, grains in crops), particularly for algorithms that derive BP as a proportion of GPP<sup>15,16</sup>. (ii) The land surface component of Earth system models currently does not take into account differences between natural and managed ecosystems which might introduce biases in BP projections. In fact, a case study based on the model ORCHIDEE<sup>17</sup> showed that taking into account a BPE difference of 8% between natural and managed ecosystems resulted in a 24% increment in BP for Europe (Supplementary Methods). (iii) Our study indicates new ways to indirectly derive BPE at regional and continental scales from maps of land use and human management. (iv) While C assimilation and BP are extensively studied, the ways to maximize BPE are less explored. However, substantial changes in yield are potentially associated with small changes in BPE. For instance, for a forest with a GPP of 1500 g C m<sup>-2</sup> y<sup>-1</sup>, an increase of 12% in BPE (Supplementary Table 7) would enhance BP by 180 g C m<sup>-2</sup> y<sup>-1</sup>, mainly in wood (Table 1). These examples show that our elucidation of BPE dynamics advances our understanding and quantification of the biomass production of terrestrial ecosystems.

199

198

183

184

185

186

187

188

189

190

191

192

193

194

195

196

197

200

201

#### References

- Luyssaert, S. *et al.* CO<sub>2</sub> balance of boreal, temperate, and tropical forests derived from a global database. *Global Change Biol.* **13**, 2509-2537 (2007).
- Vicca, S. *et al.* Fertile forests produce biomass more efficiently. *Ecol. Lett.* **15**, 520-526 (2012).
- Waring, R. H., Landsberg, J. J. & Williams, M. Net primary production of forests: a constant fraction of gross primary production? *Tree Physiol.* **18**, 129-134 (1998).
- DeLucia, E. H., Drake, J. E., Thomas, R. B. & Gonzalez-Meler, M. Forest carbon use efficiency: is respiration a constant fraction of gross primary production? *Global Change Biol.* **13**, 1157-1167 (2007).
- Forrester, D. I. Growth responses to thinning, pruning and fertiliser application in Eucalyptus plantations: A review of their production ecology and interactions. *For. Ecol. Manage.* **310**, 336-347 (2013).

- Aerts, R. & Chapin, F. S., III. in *Advances in Ecological Research* Vol. 30 (eds A. H. Fittter & D. G. Raffaelli) 1-67 (Academic Press, 2000).
- Heinemeyer, A. *et al.* Exploring the "overflow tap" theory: linking forest soil CO<sub>2</sub> fluxes and individual mycorrhizosphere components to photosynthesis. *Biogeosciences* **9**, 79-95 (2012).
- Litton, C. M., Raich, J. W. & Ryan, M. G. Carbon allocation in forest ecosystems. Global Change Biol. 13, 2089-2109 (2007).
- Treseder, K. K. A meta-analysis of mycorrhizal responses to nitrogen, phosphorus, and atmospheric CO<sub>2</sub> in field studies. *New Phytol.* **164**, 347-355 (2004).
- 223 10 Colgan, W., Carey, A. B., Trappe, J. M., Molina, R. & Thysell, D. Diversity and productivity of hypogeous fungal sporocarps in a variably thinned Douglas-fir forest. *Can. J. For. Res.* **29**, 1259-1268 (1999).
- Buée, M., Vairelles, D. & Garbaye, J. Year-round monitoring of diversity and potential metabolic activity of the ectomycorrhizal community in a beech (Fagus silvatica) forest subjected to two thinning regimes. *Mycorrhiza* **15**, 235-245 (2005).
- Mosca, E., Montecchio, L., Scattolin, L. & Garbaye, J. Enzymatic activities of three ectomycorrhizal types of Quercus robur L. in relation to tree decline and thinning. *Soil Biol. Biochem.* **39**, 2897-2904 (2007).
- Magnani, F. *et al.* The human footprint in the carbon cycle of temperate and boreal forests. *Nature* **447**, 848-850 (2007).
- Rocha, A. V. & Goulden, M. L. Why is marsh productivity so high? New insights from eddy covariance and biomass measurements in a Typha marsh. *Agric. For. Meteorol.* **149**, 159-168 (2009).
- Landsberg, J. J. & Waring, R. H. A generalised model of forest productivity using simplified concepts of radiation-use efficiency, carbon balance and partitioning. *For. Ecol. Manage.* **95**, 209-228 (1997).
- Running, S. W. & Coughlan, J. C. A general-model of forest ecosystem processes for regional applications. 1. Hydrological balance, canopy gas-exchange and primary production processes. *Ecol. Model.* **42**, 125-154 (1988).
- 243 17 Krinner, G. *et al.* A dynamic global vegetation model for studies of the coupled atmosphere-biosphere system. *Glob. Biogeochem. Cycles* **19**, GB1015 (2005).

246 Correspondence and requests for materials should be addressed to MC 247 (matteo.campioli@uantwerpen.be)

248

249

250

251

252

253

254

255

256

257

258

259

260

261

262

263

264

265

266

267

268

269

270

Acknowledgments MC, SV and JB are Postdoctoral Fellows of the Research Foundation – Flanders (FWO). SL is funded by the European Research Council (ERC) Starting grant 242564 (DOFOCO) and MFM by the Catalan Government's FI-2013 grants. We also acknowledge the ERC Synergy grant ERC-2013-SyG-610028 IMBALANCE-P. Ancillary data were provided by: G Alberti and G Delle Vedove, C Bernhofer and T Grünwald, M Jackowicz-Korczynski, B Loubet (Environment and Arable Crops Research Unit, INRA, AgroParisTech), AE Suyker, M Vadeboncoeur, D Zanotelli. This work used eddy covariance data acquired by FLUXNET and in particular by the following networks: AmeriFlux (US Department of Energy, Biological and Environmental Research, Terrestrial Carbon Program (DE-FG02-04ER63917 and DE-FG02-04ER63911)), AfriFlux, AsiaFlux, CarboAfrica, CarboEuropeIP, CarboItaly, CarboMont, ChinaFlux, Fluxnet-Canada (supported by CFCAS, NSERC, BIOCAP, Environment Canada, and NRCan), GreenGrass, KoFlux, LBA, NECC, OzFlux, TCOS-Siberia, USCCC. We acknowledge the financial support to the eddy covariance data harmonization provided by CarboEuropeIP, FAO-GTOS-TCO, iLEAPS, Max Planck Institute for Biogeochemistry, National Science Foundation, University of Tuscia, Université Laval and Environment Canada and US Department of Energy and the database development and technical support from Berkeley Water Center, Lawrence Berkeley National Laboratory, Microsoft Research eScience, Oak Ridge National Laboratory, University of California - Berkeley, University of Virginia. For sites at Kellogg Biological Station, financial support was provided by the DOE Office of Science (DE-FC02-07ER64494) and Office of Energy Efficiency and Renewable Energy (DE-AC05-76RL01830).

271	
272	Author contributions MC, SV, SL and IAJ conceived the paper; MC performed the analyses
273	and wrote the text; SL provided ORCHIDEE simulations; JB developed the multinomial
274	ordered logistic regressions and the statistics; EC, DO, DP, PFS, XW, TZ provided field data
275	or contributed to data collection from external databases and literature; all authors contributed
276	substantially to discussions and revisions.

Competing financial interests The authors declare no competing financial interests.

## Figure captions

Figure 1. Biomass production efficiency of natural and managed ecosystems. BPE (mean  $\pm$  1 s.e.m.) of (a) natural ecosystem types that can be regularly managed such as forests and grasslands; (b) natural ecosystem types that are not commonly managed such as temperate marshes, boreal peatlands and tundras, and (c) anthropogenic ecosystem types, such as croplands, that are not in a natural state but are maintained through management. Difference within forest types was significant at p<0.001 (\*\*\*), whereas difference within grassland types was significant at p<0.05 (\*). Light grey columns indicate natural ('nat.') conditions and dark grey columns managed ('man.') conditions.

Figure 2. Relationship between biomass production and gross primary production of natural and managed ecosystems. Annual values of BP and GPP with uncertainty intervals ( $s_{BPij}$  and  $s_{GPPij}$ ) reflecting measurement uncertainty and sample size (Methods) for 93 sites worldwide comprising forests, grasslands and croplands, according to the management status: managed (black, M) or natural (red, N). The slope of the linear regressions equals the biomass production efficiency,

- 298 Table 1
- 299 Title. Carbon allocation pattern in natural and managed forests as expressed by the ratio of BP
- 300 to GPP.
- 301 Footnote. Values are mean  $\pm$  1 s.e.m, in percentage; replicates (n): 12 and 19 for natural and
- 302 managed forests, respectively; notations: 'other aboveground': reproductive organs and
- 303 understory; +: 0.05<p<0.10, \*: p<0.05, and \*\*\*: p<0.001.

#### Methods

306

307

308

309

310

311

312

313

314

315

316

317

318

319

320

321

322

323

324

325

326

327

328

305

#### Dataset

Our analysis required site estimates of biomass production (BP), gross primary production (GPP), and their uncertainty, to derive the biomass production efficiency (BPE) and its uncertainty. The key rule for selecting the sites was the availability of site-specific estimates of BP and GPP. Therefore, the dataset did not include values obtained from generic algorithms (e.g. global models, remote sensing products). BP included above- and belowground growth. In most cases, BP was obtained from harvest or biometric methods (comprising empirical models as e.g. allometric relationships, root growth as function of soil conditions 18,19) and in 5% of the cases from process-based models with site-specific parameterization and/or validation against growth or biomass data. Minor gap-filling was done for BP estimates at some sites (see below). BP methodologies can be divided into broad classes according to method uncertainty (i.e. low, medium or high uncertainty<sup>1</sup>; Supplementary Table 9) related in particular to the approach to determine fine root BP (the component of ecosystem BP most difficult to assess; see Supplementary Methods) or the use of process-based models (Supplementary Table 9). However, additional tests showed that the key results of our analysis were independent of the BP methodology employed (Supplementary Table 10). GPP was mostly estimated from eddy covariance (73% of the cases) or process-based models with sitespecific parameterization and/or validation (20% of the cases). Explanation about the preference of these GPP methods instead of other approaches (e.g. GPP derived from the sum of all carbon sinks within the ecosystem such e.g. BP, autotrophic respiration, carbon transfer to mycorrhizal symbionts) is reported extensively in Supplementary Methods. Additional tests showed that the alternative use of eddy covariance- or model-based estimates of GPP did not

affect the key results of our analysis (Supplementary Table 10). Detailed information on uncertainty calculations are reported in Statistical analysis.

The integrated dataset provided BPE for 96 'golden' sites, for which BP and GPP were available from the same measuring period (53 forests, 14 grasslands, 24 croplands and 5 wetlands) and 35 additional natural sites for which BP and GPP were both available but not for the same measuring period (16 forests, 6 grasslands, 8 wetlands, 5 tundra). Wetlands were divided into marshes (herbaceous-dominated vegetation of the temperate zone mainly affected by flooding from river, sea or irrigation; 6 in total) and peatlands (ombotrophic or minerotrophic inland boreal ecosystems rich in herbs, shrubs or mosses; 7 in total). An excerpt from the study dataset is shown in Supplementary Table 1 and the geographical distribution of the sites in Supplementary Fig. 1. The key data used in the analysis are provided in Supplementary Data.

Ancillary data such as vegetation characteristics, climate, environmental conditions and anthropogenic impacts were needed for each site to determine the possible effect of these factors on BPE. Such information was retrieved mostly from the literature, open-access databases<sup>1,20-27</sup> or modelling<sup>28</sup> (Supplementary Table 2). For N deposition, data for Western Europe and the conterminous USA were retrieved from interpolated gridded maps based on ground observations<sup>25</sup>, whereas simulated values were used for the rest of the world<sup>22,23</sup>.

### Management classification

The sites were divided into two categories: natural and managed. Natural sites are defined as those characterized by none or low-to-moderate human impact, whereas managed sites are heavily affected by human activity. We defined 'low-to-moderate human impacts' as human activities that largely reproduce naturally occurring processes e.g. low grazing, occasional fire in grasslands, forest regeneration. On the other hand, we considered sites 'heavily affected by

human activity' to be those with impacts that would not occur in nature e.g. intense fertilization of poor soils, sowing of cropland monocultures, thinning of healthy trees. The classification was straightforward for marshes, peatlands and tundras (pristine or with minimal human impact except in two managed wetlands) and for croplands (inherently managed) (Supplementary Table 3). For forests and grasslands, the classification included sub-categories for both the natural and managed classes (Supplementary Table 3). For instance, for forests, we considered as natural the following types of forests: (i) old-growth with minimal disturbance, (ii) natural succession due to fire/windthrow and at least 10 years after the disturbance, (iii) unmanaged or with low human impact (e.g. understory grazing) in the 50 years before measurement and (iv) planted forests without any intervention after planting and at least 10 years old at the time of measurement. On the other hand, we considered as managed forests: (i) forests with thinning/harvest in the 50 years before measurement, (ii) newly (<10 years old) established plantations, (iii) forests fertilized in the 25 years before measurement, or (iv) forests managed for fruit/rubber production at time of measurement. Similar sub-categories were defined for grasslands (Supplementary Table 3).

We tested the validity of our approach by comparing our binary management classification to a more complex three-level classification According to the latter approach, we considered 'pristine natural' the sites that were pristine or with minimal impacts and 'semi-natural' the sites with low-moderate human impacts (these classes were considered jointly in the binary classification as 'natural'). For forests, for instance, we considered as semi-natural the forests that were: (i) unmanaged or with low human impact (e.g. understory grazing) in the 50 years before measurement and (ii) planted forests without any intervention after planting and at least 10 years old (see above). The statistics of this additional test clearly showed that (i) BPE of pristine natural and semi-natural forests did not differ and that (ii) the BPE difference between pristine natural and semi-natural forests was considerably lower than

the difference between semi-natural and managed forests (Supplementary Fig. 4, Supplementary Table 11). This confirmed that our standard binary classification is sound and that our key result about the impact of management on BPE is robust. In addition, this exercise revealed that the introduction of more levels in the management classification would not be advantageous, but rather would make the entire statistical analysis more complex and less robust. This was evident for grasslands, for which the three-level classification did not alter the BPE pattern but substantially reduced the statistical power because of the smaller sample size and associated higher uncertainties (Supplementary Fig. 4, Supplementary Table 11).

## Gap-filling

Some of the selected sites lacked BP measurements of minor ecosystem biomass components (e.g. nonvascular plants, understory) or were affected by minor systematic measurement biases (e.g. neglecting litterfall decomposition in tropical forests). These missing BP portions were gap-filled for completeness in analogy to Vicca et al 2012<sup>2</sup>.

Production of reproductive organs in forests. When missing, this BP component was derived from a relationship between reproductive BP versus aboveground BP<sup>2</sup> derived from the Global Forest Database<sup>1</sup>.

Leaf biomass production in tropical forests. Estimates of leaf BP in tropical forests are systematically underestimated because of within-canopy decomposition of leaf litter during the collection period. We estimated this missing portion of BP as 12% of total foliage production<sup>2</sup>.

Understory biomass production in forests. BP due to understory vegetation is significant for boreal forests and thus boreal forests lacking this BP component were not considered in our analysis<sup>2</sup>. However, the contribution of understory BP to total ecosystem BP

is more limited for temperate and tropical forests<sup>2</sup>. Thus, we did not discard temperate and tropical forests lacking understory BP but gap-filled this missing BP component as done in previous studies<sup>2</sup>. In particular, understory BP was estimated as a fixed ratio of the forest tree BP: 0.043 for temperate and 0.073 for tropical forests<sup>2</sup>.

Nonvascular biomass production in tundra. Missing nonvascular BP was derived from a nonvascular productivity ratio (BP-to-biomass ratio, the portion of biomass renewed every year). This ratio was calculated for wet (0.50 years<sup>-1</sup>) and mesic tundra (0.42 years<sup>-1</sup>) as the average of six observations for each tundra type (Supplementary Table 12).

Shrub biomass production due to stem secondary growth in peatland. Missing BP due to unaccounted shrub secondary growth (i.e. increase in stem/branch diameter) was estimated to be 29% of the shrub aboveground primary growth (i.e. BP due to current-year leaves and stem/branches) from data for subarctic shrubs<sup>29</sup>.

The gap-filling concerned 31 forests of the 96 golden sites and 17 sites (14 forests, two tundra and one peatland) of the additional 35 natural sites. For 69% of the cases, the gap-filled BP differed by less than 5% than the original BP; for 13% of the cases the gap-filled and original BP differed by 5-10%, whereas for 17% of the cases this difference was 10-15%. Herbivory was not taken into account because it was negligible (e.g. for forests<sup>2</sup>) or because BP measurements were from experiments that excluded large herbivores (e.g. for all grasslands examined).

The gap-filling procedure avoided small secondary biases in the analysis but did not alter the primary results (Supplementary Table 13). Overall, original BPE of managed and natural forests (the ecosystem type most affected by gap-filling) was  $0.52 \pm 0.03$  and  $0.39 \pm 0.02$  (mean  $\pm$  s.e.m.), respectively, which was less than 2% smaller than gap-filled BPE (Supplementary Table 7).

430

## Statistical analysis

431

432

433

434

435

436

437

438

439

440

441

442

443

444

445

446

447

448

449

450

451

452

453

Analysis overview and dataset

Our study consisted of five analyses, using different sub-sets of our database. (1) First, we analyzed all natural sites (n=75; managed sites were not considered in this analysis) to test whether BPE is driven by natural variation in site fertility. In particular, we tested whether BPE differs among ecosystem types and sites of contrasting fertility. (2) Second, we analyzed the relationship between fertility and management in forests to verify their correlation and disentangle (i) the impact of management on fertility from (ii) the fertility status not related to management. This analysis was performed on 53 managed and natural forests for which BP and GPP were measured during the same period. We focused this analysis only on forests because they are the ecosystem type best represented in our dataset and allow direct comparison with previous studies. (3) Third, the relative importance of fertility, management and N deposition as controllers of BPE was compared to the importance of other possible BPE drivers. This analysis was performed on the same forest dataset considered in the second analysis after disentangling the effect of fertility and management. (4) Fourth, we compared the BPE of key natural and managed ecosystem types (grasslands, forests and croplands) that typically share similar climatic and environmental characteristics and are regularly converted into one another. Only sites with BPE obtained from BP and GPP measured during the same period were used (n=93). (5) Five, we studied the impact of the potential drivers of BPE in all natural ecosystems (n=75; this analysis did not include the managed sites).

For the analyses 1 and 5, we considered not only the sites for which BP and GPP were measured during the same period but also sites with BP and GPP measured during different (or only partially overlapping) periods (35 out of the 75 sites) to investigate a large set of

ecosystem types (e.g. from forest to tundra) and environmental conditions (e.g. climate from tropical to polar, soil from waterlogged to very dry). For sites without management operations (and mostly at mature-old stage) the temporal mismatch in BP and GPP was less crucial, dampened at several sites by multi-year measurements (we used averages of BP and GPP for multi-year observations) and, most importantly, comparative tests revealed that the results of the analyses did not differ when all sites or only sites with temporal match in BP and GPP were considered (e.g. Supplementary Table 14).

Relationship between fertility and management

Site fertility and site management are highly correlated factors that are both potentially crucial for BPE. For this study, we wanted to separate both drivers to test for BPE responses to (1) the fertility status induced by management and (ii) the fertility status unrelated to management. To disentangle both effects, we applied an approach commonly used to deal with multicollinearity<sup>30</sup>: the observed fertility status was modeled as a function of management and the residuals from this model were used as explanatory variables of BPE (instead of the original fertility status). Hence, the residuals reflect the information on fertility not explained by management, which we termed 'unexplained natural fertility'. Initially, the model also included N deposition as an additional covariate, but we removed it in the final model as the relationship between N deposition and fertility was weak (see Main text).

A multinomial ordered logistic regression model (or 'proportional odds logistic regression model'<sup>30</sup>) was fitted with fertility as outcome (ordinal categorical variable with category high, H, medium, M, and low, L) and management (yes/no) as covariate. The model estimates the log odds of falling into or below a fertility category as a function of management:

- 479 Logit P(fertility=L)=intercept<sub>L</sub> +  $\beta_L \times$  management (1)
- 480 Logit P(fertility<M)=intercept<sub>M</sub> +  $\beta_M \times$  management (2)

where intercept<sub>L</sub> and intercept<sub>M</sub> were -2.01 and -0.511, respectively, and  $\beta_L$  and  $\beta_M$  were 2.84 and -0.0488, respectively. In other words, this model estimates the possible fertility distribution of each site according to its management status (given its management status, the probability to be H, M or L). Also three sets of residuals were obtained for each site, which reflect the deviation of the fertility status of the site from the distribution estimated by the model. The independence of these three residuals on management (unlike the original fertility variable) was proven with t-tests (all p-values > 0.05).

### BPE drivers

The relationships between BPE and its potential drivers were explored with three statistical approaches: univariate analysis, multiple linear regressions and Random Forest, which are described below. We used the following predictors: management status, observed natural fertility, climate zone, ecosystem type, growth form (five categorical variables) and N deposition, unexplained natural fertility (the three model residuals described above), soil available water content, annual precipitation and dry months per year (seven continuous variables) (Supplementary Table 2). All analyses were performed with R<sup>31</sup>.

Univariate analysis. This analysis tested the significance of the relationships between single predictors and BPE. For continuous variables, this was done with single linear regressions, whereas for categorical variables we used one-way ANOVAs with post-hoc Tukey's HSD test. Normality of residuals was tested with Shapiro-Wilks' test and the assumption of homoscedasticity with Levene's test (for ANOVAs) or Breusch-Pagan test (for regressions). For the few cases for which these conditions were not met, data were

transformed (e.g. log(x), 1/x or  $x^2$ ) or treated with alternative methods (Kruskal-Wallis test for non-normality and applications of White method for heteroskedasticity<sup>32</sup>).

Multiple linear regressions. This method allows a comparison of the effect of the potential BPE predictors considering them all together. Whenever a given predictor was significant in the univariate analysis, but not in the multiple linear regressions, this indicated a lower importance of that predictor as compared to other predictors. In practice, we opted for backward stepwise regressions. Accordingly, the best BPE model was determined by starting from the model with all variables and successively removing the least important. The selection was done by comparing the new model (without the removed variable) with the original model (with the original variable) using Likelihood Ratio and Akaike Information Criterion (AIC). In practice, the new model was not accepted if the Likelihood Ratio was significant (p<0.05) or the AIC increased. Stepwise multiple linear regression was a suitable methodology for our analysis, because it can be applied with both continuous and categorical variables. However, all factors of categorical variables need to be taken into consideration by introducing dummy variables. Prerequisites (or alternatives) for applying linear regressions (e.g. residuals normality and homoscedasticity) were tested as described above for Univariate analysis.

Random Forest. We used this partitioning method to produce a large ensemble of regression trees considering always our complete BPE dataset but random subsets of predictor variables<sup>33</sup>. This means that (in contrast to multiple linear regressions) Random Forest accounts also for non-linear relationships and interactions, and evaluates each predictor variable (even the least important or redundant), providing a ranking of the predictors' importance. However, this analysis does not assign a significance label (contrary to linear regressions analysis). The importance of a given variable is instead indicated by the mean decrease in accuracy (or increase in mean squared error, %IncMSE) of model predictions

when the value of that given variable was changed (permuted within the dataset)<sup>33</sup>. The more important the variable, the larger the difference between original predictions and new predictions, and the larger the %IncMSE. We used the standard Random Forest algorithm<sup>34</sup> setting a large number of trees (50000) to obtain stable results.

#### Confounding factors

The response of BPE to N deposition and variables related to the water status (soil available water content, precipitation, dry months per year) could have been confounded by fertilization and irrigation / exceptional soil water conditions, respectively, at some sites. To check for the relevance of confounding factors, the analyses comprising N deposition and the variables related to the water status were performed both on the entire dataset and on a subset that excluded sites with fertilization, irrigation, occasional flooding, minerotrophic conditions and permafrost. Overall, the impact of these sites was negligible (Supplementary Table 15) and therefore they were not removed in the final analyses. Through the analysis, filtering for outliers was minimal and we removed only four sites with unrealistic BPE (0.84-0.94).

# *Uncertainty*

- The BP uncertainty for site i ( $s_{\text{BPij}}$ ) depended on (i) a typical range of uncertainty ( $p_{\text{BPi}}$ ) based
- on ecosystem type, (ii) the experimental methodology *j* through a method-specific uncertainty
- reduction factor  $(RF_{BPj})$  and (iii) the length of the measurement period in years  $(l_{BPij})^1$ :

549 
$$s_{BPij} = \frac{(p_{BPi} \times RF_{BPj})}{(l_{BPij})^{0.5}} (3)$$

- In case BP needed to be gap-filled (see above), the uncertainty of the original BP estimate
- $(s_{\text{BPij original}})$  was increased by a factor equivalent to 100% of the gap-filling amount<sup>2</sup>:

552 
$$s_{BPij \ gapfilled} = \left( \left( s_{BPij \ original} \right)^2 + \left( gapfilling \right)^2 \right)^{0.5} (4)$$

where  $s_{\text{BPij gapfilled}}$  is the uncertainty of the gap-filled BP estimate. The uncertainty of GPP

554  $(s_{\text{GPPij}})$  was calculated in the same way as  $s_{\text{BPij}}$ :

$$S_{GPPij} = \frac{(p_{GPPi} \times RF_{GPPj})}{(l_{GPPij})^{0.5}} (5)$$

where  $p_{\text{GPPi}}$  is the typical range of GPP uncertainty,  $RF_{\text{BPi}}$  the uncertainty reduction factor

dependent on the experimental methodology j and  $l_{GPPij}$  the length of the measurement period

in years. The uncertainty of BPE ( $s_{\text{BPE}ii}$ ) was calculated through error propagation:

$$559 s_{BPEij} = \left( \left( \frac{s_{BPij}}{BP_{ij}} \right)^2 + \left( \frac{s_{GPPij}}{GPP_{ij}} \right)^2 \right)^{0.5}$$
(6)

where  $BP_{ij}$  and  $GPP_{ij}$  are values of BP and GPP, respectively, for site i and method j. Values

of  $RF_{\rm BPj}$  and  $RF_{\rm GPPj}$  were determined following Luyssaert et al 2007<sup>1</sup> and were reported in

Supplementary Table 9. For forest ecosystems, values of  $p_{BPi}$  and  $p_{GPPi}$  were from Luyssaert et

al 2007<sup>1</sup>, whereas for non-forest ecosystems they were derived from the difference between

the ninth and first decile of BP and GPP samples from ca. 20 to 110 sites according to

ecosystem type (Supplementary Table 16).

# 566

563

564

558

567

568

## **References only in Methods**

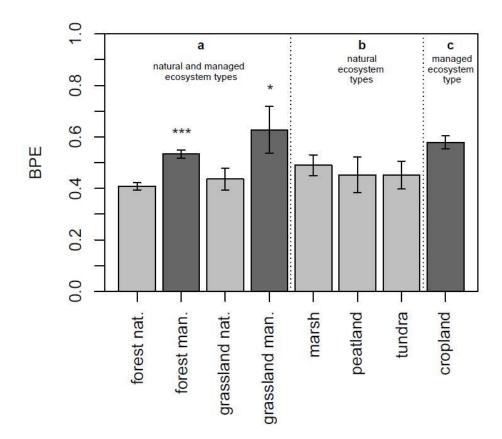
- Aber, J. D., Melillo, J. M., Nadelhoffer, K. J., McClaugherty, C. A. & Pastor, J. Fine roots turnover in forest ecosystems in relation to quantity and form of nitrogen availability- a comparison of 2 methods. *Oecologia* **66**, 317-321 (1985).
- Baret, F., Olioso, A. & Luciani, J. L. Root biomass fraction as a function of growth degree days in wheat. *Plant Soil* **140**, 137-144 (1992).
- 574 20 Batjes, N. H. *ISRIC-WISE global data set of derived soil properties on a 0.5 by 0.5* 575 *degree grid (Version 3.0).* Vol. Report 2005/08 (ISRIC - World Soil Information, Wageningen (with data set), 2005).
- 577 21 Hijmans, R. J., Cameron, S. E., Parra, J. L., Jones, P. G. & Jarvis, A. Very high resolution interpolated climate surfaces for global land areas. *Int. J. Climatol.* **25**, 1965-1978 (2005).
- Dentener, F. J. Global maps of atmospheric nitrogen deposition, 1860, 1993, and 2050. (Data set. Available on-line [http://www.daac.ornl.gov] from Oak Ridge National Laboratory Distributed Active Archive Center, 2006).
- 583 23 Galloway, J. N. *et al.* Nitrogen cycles: past, present, and future. *Biogeochemistry* **70**, 153-226 (2004).

- Holland, E. A., Braswell, B. H., Sulzman, J. & Lamarque, J. F. Nitrogen deposition onto the United States and western Europe: Synthesis of observations and models. *Ecol. Appl.* **15**, 38-57 (2005).
- Holland, E. A., Braswell, B. H., Sulzman, J. M. & Lamarque, J. F. *Nitrogen deposition* onto the United States and Western Europe. (Data set. Available on-line [http://www.daac.ornl.gov] from Oak Ridge National Laboratory Distributed Active Archive Center, 2005).
- Harris, I., Jones, P. D., Osborn, T. J. & Lister, D. H. Updated high-resolution grids of monthly climatic observations - the CRU TS3.10 Dataset. *Int. J. Climatol.* **34**, 623-642 (2014).
- Jones, P. D. & Harris, I. CRU TS3.10: Climatic Research Unit (CRU) Time-Series (TS)
  version 3.10 of high resolution gridded data of month-by-month variation in climate
  (Jan. 1901 Dec. 2009). University of East Anglia Climatic Research Unit. (NCAS
  British Atmospheric Data Centre, 2013).
- Schaap, M. G., Leij, F. J. & van Genuchten, M. T. ROSETTA: a computer program for estimating soil hydraulic parameters with hierarchical pedotransfer functions. *J. Hydrol.* **251**, 163-176 (2001).
- 602 29 Campioli, M., Leblans, N. & Michelsen, A. Stem secondary growth of tundra shrubs: impact of environmental factors and relationships with apical growth. *Arct. Antarct. Alp. Res.* **44**, 16-25 (2012).
- 605 30 Agresti, A. An Introduction to Categorical Data Analysis. (John Wiley & Sons, Inc., 2002).
- Team, R. D. C. *R: A language and environment for statistical computing.* (R Foundation for Statistical Computing, 2013).
- Zeileis, A. Object-oriented computation of sandwich estimators. *J. Stat. Softw.* **16**, 9 (2006).
- 611 33 Breiman, L. Random forests. *Mach. Learn.* **45**, 5-32 (2001).

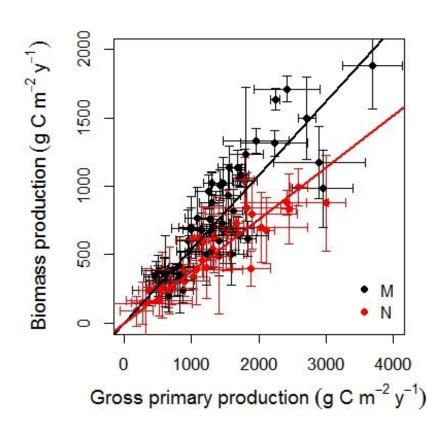
615

612 34 Liaw, A. & Wiener, M. Classification and regression by Random Forest. *R news* **2**, 5 (2002).

Figure 1



620 Figure 2



624 Table 1

BP-to-GPP ratio	N (n=12)	M (n=19)	p diff N-M
leaves	10±1%	10±1%	0.91
wood	11±1%	$24 \pm 3\%$	0.00019***
other aboveground	6±2%	7±3%	0.61
fine roots	12±2%	8±2%	0.083 +
coarse roots	3±1%	$4\pm1\%$	0.29
whole ecosystem (BPE)	41+2%	53+3%	0.020*